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SIMULATING RADAR SIGNALS FOR DETECTION PERFORMANCE EVALUATION.(U)

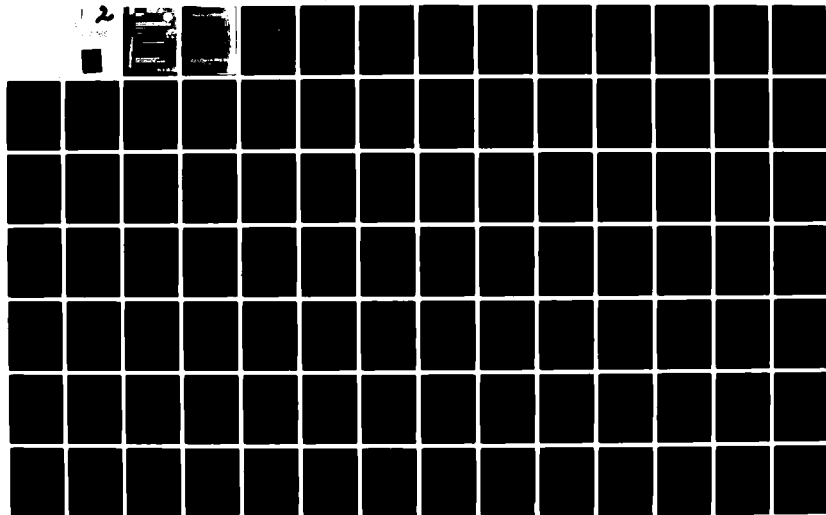
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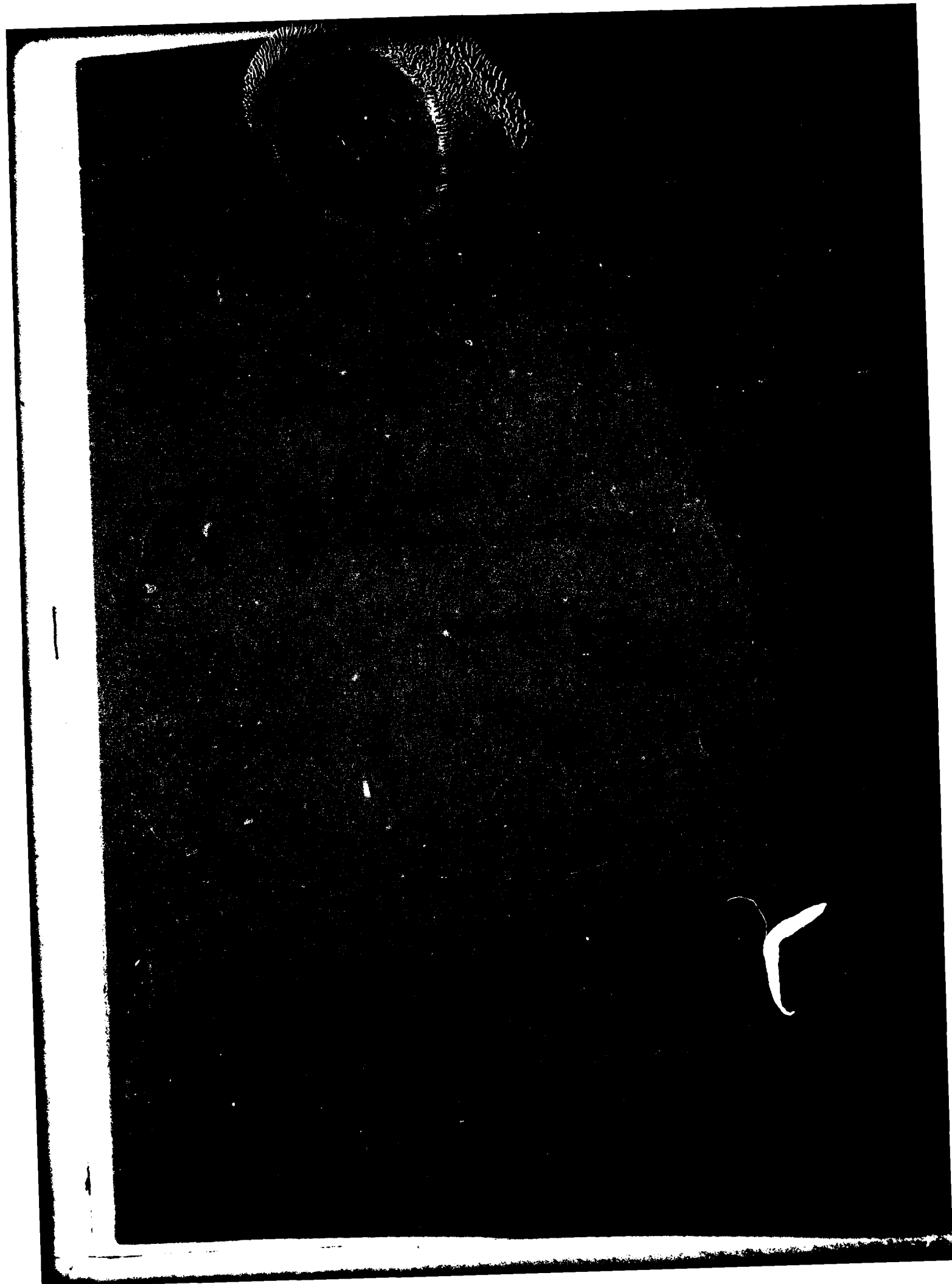
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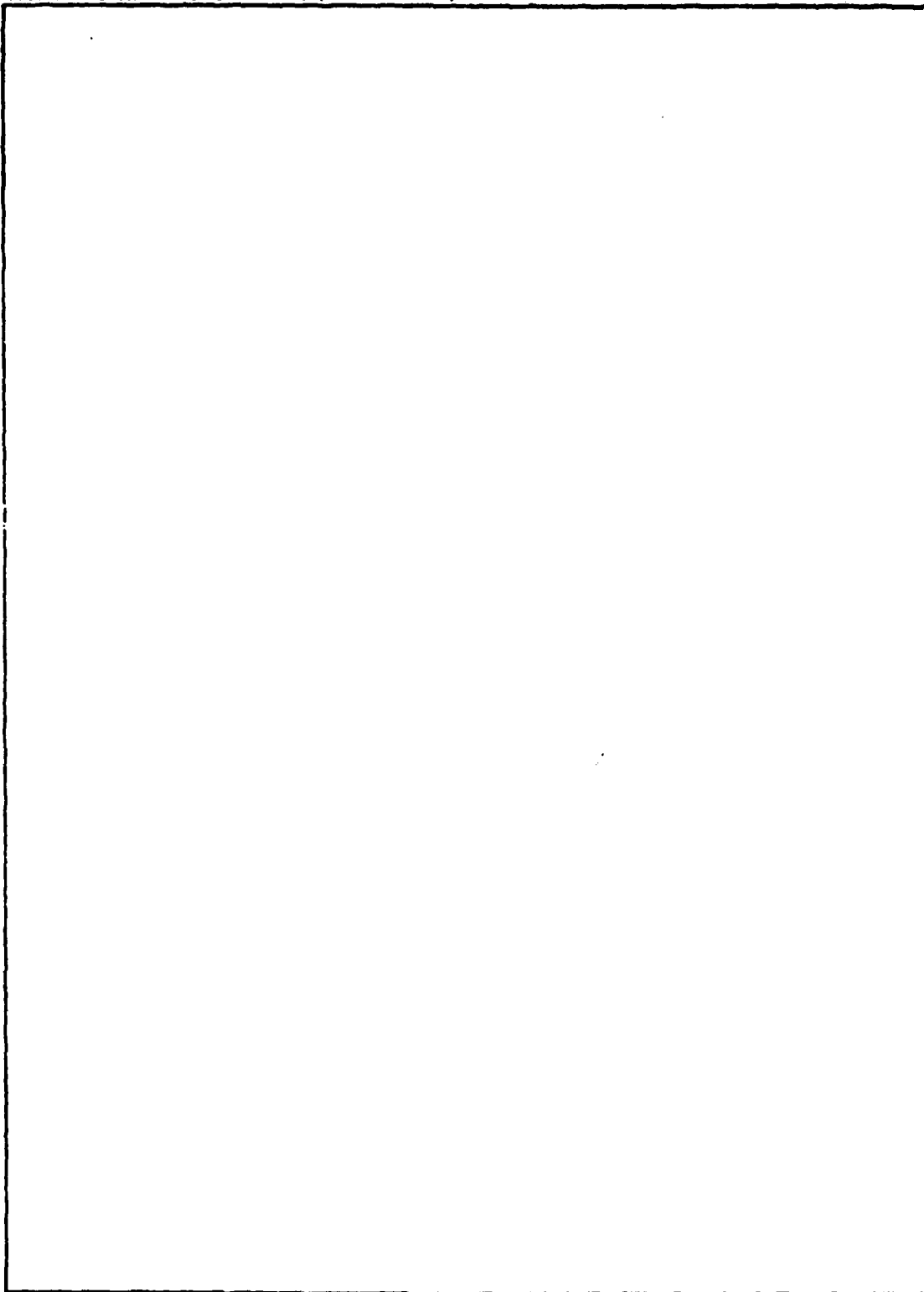
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EVALUATION

The objective of the study was to develop a method of determining the detection performance of ground-based radars against airborne targets under a wide variety of environment and system conditions.

This effort involved the development of a novel computationally efficient method for simulating ground-based coherent radar system performance subjected to target and clutter signals which are partially correlated.

The shortcoming of conventional simulation techniques used to compute detection performance is that an extremely large number of statistical replications are required to establish false alarm performance. This work advanced the art by adapting the technique of importance sampling which significantly reduces the required computer time, even when simulating nonlinear systems. In most cases it is possible to simulate false alarm probabilities as low as 10^{-8} with only about 10^3 to 10^4 replications. A report, entitled "Importance Sampling Applied to Radar False-Alarms," summarizing results and application examples has been accepted for publication in the IEEE Transactions of the AES.


J. LEON POTRIER
Contract Monitor

1. INTRODUCTION AND SUMMARY

The objective of the study described in this report was to develop a method of determining detection performance for a ground-based radar against an airborne target under a wide variety of environment and system conditions. This objective was met by constructing a digital simulation of the radar signals as they are transmitted, reflected from the scatterers representing targets and clutter, and received by and processed in the radar receiver. Realistic ground, rain, and chaff clutter environments are modeled, as well as the processing features of modern radars such as moving target indication (MTI), both coherent and noncoherent integration, constant false alarm rate (CFAR) processing, and non-linear operations. The detection performance is determined by Monte Carlo sampling techniques, in which the target and clutter scattering models are described statistically, as well as the location of the target with respect to the center of the antenna beam, range gate, and Doppler filter.

The shortcoming of conventional sampling techniques applied to detection performance in radar is the extremely large number of statistical replications required to establish false alarm performance. In order to overcome this problem a considerable effort was expended in developing importance sampling techniques that could be applied to the wide variety of signals in radar, including the non-Gaussian and non-Rayleigh signals characteristic of clutter. The results of this effort are described in Section 2. In most cases it is possible to simulate false alarm rates as low as 10^{-8} with only about 10^3 to 10^4 replications of the experiment. Techniques for handling mixed statistics as well as CFAR are also described.

It is well known that antenna motion during the coherent processing time of the radar causes the clutter to be amplitude modulated; the effect is a broadening of the clutter spectrum that would be observed if there were no antenna motion. It is straightforward to compute the resultant spectrum if the clutter is spatially homogeneous, especially if the antenna pattern and original clutter spectrum are Gaussian shaped. In Section 3 we extend the analysis to nonhomogeneous clutter and arbitrary beamshapes. While the signals are nonstationary, the resulting algorithms are in a form that is amenable to efficient digital simulation.

In Section 4 a general procedure is derived for generating correlated random signal sequences that are characteristic of clutter in a ground-based radar. The

procedure is extremely efficient as it is based on fast Fourier transform (FFT) combined with interpolation. It can be applied to any number of signal samples and the shape of the power spectral density is arbitrary.

The simulation program that was developed to determine detection performance is described in Section 5, and a Fortran listing is given in Appendix A.

2. IMPORTANCE SAMPLING

Importance sampling is a technique that can be applied to the simulation of low-probability events without incurring the computation costs usually associated with such simulations. With importance sampling one can modify the probability distribution of the underlying random process in order to make the low-probability events (false alarms) occur more frequently. The desired probabilities at the output of the process are then found by weighting each event by a factor that is a function of only the state of the input; this factor is independent of the process itself [1-6].

The basic principle of importance sampling is straightforward as described above. However, it is not so well known just how the technique can be made to work in a particular application. For example, what if there are multiple random inputs to some processor where the inputs might belong to different statistical processes? Or if the processor is nonlinear? Or if there is not a unique relationship between the input and output of the processor? These and other issues will be addressed in this report. We will concentrate on applications to the simulation of signals in radar and communication systems in order to limit the scope of the study. We begin with a tutorial discussion of both conventional and importance sampling.

2.1 CONVENTIONAL SAMPLING THEORY

Let us designate the output of a statistical process as y . We wish to estimate the probability density function of this process, $p(y)$, or its cumulative distribution function, $P(y)$, with a finite set of observations. The conventional procedure is to sort the output samples into preselected bins that cover the range of interest in the variable y . This operation of sorting results in a histogram, which can be integrated to form the sample distribution function.

For the purpose of examining the upper tail of the distribution function it is more convenient to work with the complement to $P(y)$, namely

$$Q(y) = 1 - P(y) = \int_y^{\infty} p(\xi) d\xi \quad (1)$$

In estimating $Q(y)$ with a finite set of observations, we can define the operation applied to each sample as

$$\begin{aligned} D_Y(y) &= 1 & , & & y \geq Y \\ &= 0 & , & & y < Y \end{aligned} \quad (2)$$

where Y is a preset threshold. In practice, this operation will be applied to many values of Y simultaneously, for each sample of y , in the prior step of computing the histogram. We note that all samples receive the same weight.

The result of applying the operation in (2) to each sample of y is also a statistical process. The mean value of (2) is

$$\begin{aligned} \overline{D_Y(y)} &= \int D_Y(y) p(y) dy \\ &= \int_Y^{\infty} p(y) dy = Q(Y) \end{aligned} \quad (3)$$

and the second moment is

$$\begin{aligned} \overline{D_Y^2(y)} &= \int D_Y^2(y) p(y) dy \\ &= \int_Y^{\infty} p(y) dy = Q(Y) \end{aligned} \quad (4)$$

The variance is given by

$$\begin{aligned} \text{var}[D_Y(y)] &= \overline{D_Y^2(y)} - [\overline{D_Y(y)}]^2 \\ &= Q(Y) - Q^2(Y) = Q(Y)P(Y) \end{aligned} \quad (5)$$

In the following analysis we will be interested in the upper tail of the distribution where $P(Y) \approx 1$, so we can essentially assume that

$$\text{var } [D_Y(y)] = Q(Y) \quad . \quad (6)$$

In order to estimate $Q(Y)$ the operation in (2) will be repeated for N samples of y and the estimate will be formed as

$$\hat{Q}(Y) = \frac{1}{N} \sum_{i=1}^N D_Y(y_i) \quad , \quad (7)$$

where $\{y_i\}$ is the set of N observations (statistical samples), which we assume to be independent. It follows from (3) and (7) that $\overline{\hat{Q}(Y)} = Q(Y)$, which means that (7) is an unbiased estimate. The variance of (7) is given by

$$\begin{aligned} \text{var } [\hat{Q}(Y)] &= \frac{1}{N} Q(Y)P(Y) \\ &= \frac{1}{N} Q(Y) \quad . \end{aligned} \quad (8)$$

In order to estimate $Q(Y)$ with high precision, the standard deviation of the estimate must be small compared with the mean value. In other words, $NQ(Y) \gg 1$. If, for example, $Q(Y) = 10^{-6}$ then N must be at least as large as 10^6 in order to achieve any precision at all in the estimate. This requirement for an impractically large number of samples is the dilemma faced when one applies conventional sampling techniques to the estimate of low-probability events.

Importance sampling will be a solution to this dilemma, but before we jump to that subject let us expand our discussion to include a description of the process itself. As sketched in Figure 1, the input to some processor will be a random variable x with a known probability density function $p(x)$. The output is y , for which we wish to estimate $Q(y)$, the complement to the distribution function, with a finite set of observations. The estimate will be formed by taking the average value of the observable z , which is the operation $D_Y(y)$ applied to y . The transfer function of the processor will be designated by $y = F(x)$, which implies that a given value of x is mapped into a unique value of y . Since each x results in a particular value of z we can write

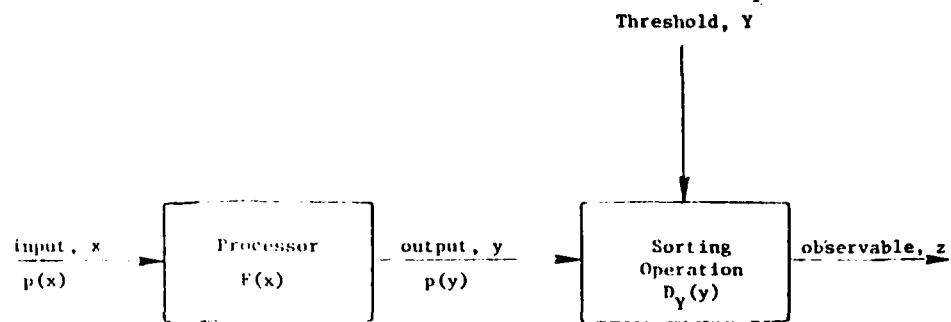


Figure 1. Sketch of Operations Performed with Conventional Sampling

$$z = D_Y[F(x)] \quad (9)$$

2.2 IMPORTANCE SAMPLING THEORY

The principle of importance sampling is to distort or modify the input random process $\{x\}$ in order to make the original low-probability events occur more frequently. This action will be compensated by weighting the event by a factor that is a function of only the particular value of x on input $\{1,2,3\}$. The functional flow is sketched in Figure 2 where

$$z_m = D_Y[F(x)]w(x) \quad (10)$$

The weight is designated as $w(x)$ and the modified probability density function on input is designated as $p_m(x)$. In order for z_m to be an unbiased estimate of $Q(Y)$ we must have $\bar{z}_m = \bar{z}$, the latter quantity being the mean value of (9), so that

$$\int D_Y[F(x)]w(x)p_m(x)dx = \int D_Y[F(x)]p(x)dx \quad (11)$$

thus

$$w(x) = p(x)/p_m(x) \quad (12)$$

For each replication of the experiment a specific value of x will be generated. With the use of (12) we can then compute the weight on the basis of the ratio of known input probability density functions.

So far we have been working with a one-dimensional process. Actually the process could be multi-dimensional. We will assume that there is still a single output, y , so that we can write

$$z = D_Y[F(x_1, \dots, x_k)] \quad (13)$$

and

$$z_m = D_Y[F(x_1, \dots, x_k)]w(x_1, \dots, x_k) \quad (14)$$

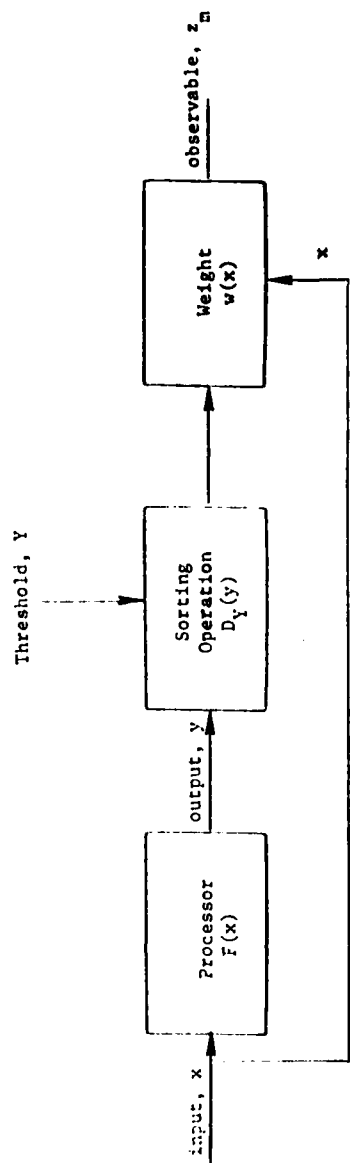


Figure 2. Sketch of Operations Performed with Importance Sampling

And since $\bar{z}_m = z$, we have

$$w(x_1, \dots, x_K) = \frac{p(x_1, \dots, x_K)}{p_m(x_1, \dots, x_K)} \quad (15)$$

where the right side of (15) is the ratio of two joint probability density functions. If the input samples are independent and belong to the same distribution function,

$$w(x_1, \dots, x_K) = \prod_{k=1}^K p(x_k)/p_m(x_k) \quad (16)$$

2.3 SPECIFIC EXAMPLES

In order to demonstrate the utility and power of importance sampling, several examples will be given. First, we will examine linear processors where the output statistics are known, and then we will analyze some non-linear processors.

Example 1: Exponential Distribution

Let x be exponentially distributed with $y = F(x) = x$. We can define

$$\begin{aligned} p(x) &= (1/\bar{x})e^{-x/\bar{x}}, & x &\geq 0 \\ &= 0, & x &< 0 \end{aligned} \quad (17)$$

where \bar{x} is the mean value of the original distribution. Upon integration we obtain

$$Q(x) = e^{-x/\bar{x}} \quad (18)$$

for $x \geq 0$. For importance sampling we will modify the input distribution by changing (increasing) the mean value as

$$p_m(x) = \frac{1}{\bar{x}_m} e^{-x/\bar{x}_m}, \quad x \geq 0$$

$$= 0, \quad x < 0. \quad (19)$$

The weight is now given by

$$w(x) = \frac{p(x)}{p_m(x)} = \frac{\bar{x}_m}{x} e^{-(1/\bar{x} - 1/\bar{x}_m)x}, \quad (20)$$

which is a simple calculation that will be performed for each sample on input.

For the case when $y = F(x) = x$ as we have here, the second moment of z_m can be evaluated easily. For a single observation

$$\begin{aligned} z_m^2 &= \int_0^\infty D_Y^2(x) w^2(x) p_m(x) dx \\ &= \int_Y^\infty w(x) p(x) dx \end{aligned} \quad (21)$$

For the exponential distribution we can substitute (17) and (20) to obtain

$$z_m^2 = \frac{\bar{x}_m}{x(2 - x/\bar{x}_m)} e^{-(2/\bar{x} - 1/\bar{x}_m)Y}. \quad (22)$$

Since $\bar{x} = Q(Y) = e^{-Y/\bar{x}}$, the variance for a single observation is given by

$$\text{var}[z_m] = \left[\frac{\bar{x}_m}{x(2 - x/\bar{x}_m)} e^{Y/\bar{x}_m} - 1 \right] e^{-2Y/\bar{x}}. \quad (23)$$

For N independent observations in estimating the distribution function, the variance will be reduced by a factor of N over that of a single observation. Thus we can write the following ratio for N independent observations:

$$\frac{\text{var}[z_m]}{(\bar{z}_m)^2} = \frac{1}{N} \left[\frac{\bar{x}_m}{\bar{x}(2 - \bar{x}/\bar{x}_m)} e^{Y/\bar{x}_m} - 1 \right] \quad (24)$$

When we work with the expression in (24) it is convenient to simplify it slightly by anticipating the result $\bar{x}_m \gg \bar{x}$. Thus

$$\frac{\text{var}[z_m]}{(\bar{z}_m)^2} = \frac{1}{N} \left[\frac{\bar{x}_m}{2\bar{x}} e^{Y/\bar{x}_m} - 1 \right] \quad (25)$$

Now we can ask the question as to what the optimum value of \bar{x}_m is. If we differentiate (25) with respect to \bar{x}_m and set the result equal to zero we obtain $\bar{x}_m = Y$ as the solution. At the optimum,

$$\frac{\text{var}[z_m]}{(\bar{z}_m)^2} = \frac{1}{N} \left[\frac{e}{2} \frac{Y}{\bar{x}} - 1 \right] \quad (26)$$

For illustration, suppose we are interested in estimating $Q(Y)$ in the neighborhood of $Q(Y) = 10^{-6}$. With (18) we obtain $\bar{x}_m = Y = 13.8\bar{x}$ (which means that the approximation made in obtaining (25) was indeed valid) and (26) reduces to

$$\frac{\text{var}[z_m]}{(\bar{z}_m)^2} = 17.8/N \quad (27)$$

In Section 2 with conventional sampling techniques this ratio was $10^6/N$. We have thus reduced the number of samples required to obtain a specified precision by a factor of over 50,000. With $N = 1000$ samples the relative standard deviation in the estimate of $Q(Y)$ will be 13.3% with importance sampling, an acceptable error for practically any application.

In Figure 3 we compare the experimental distribution function with $N = 1000$

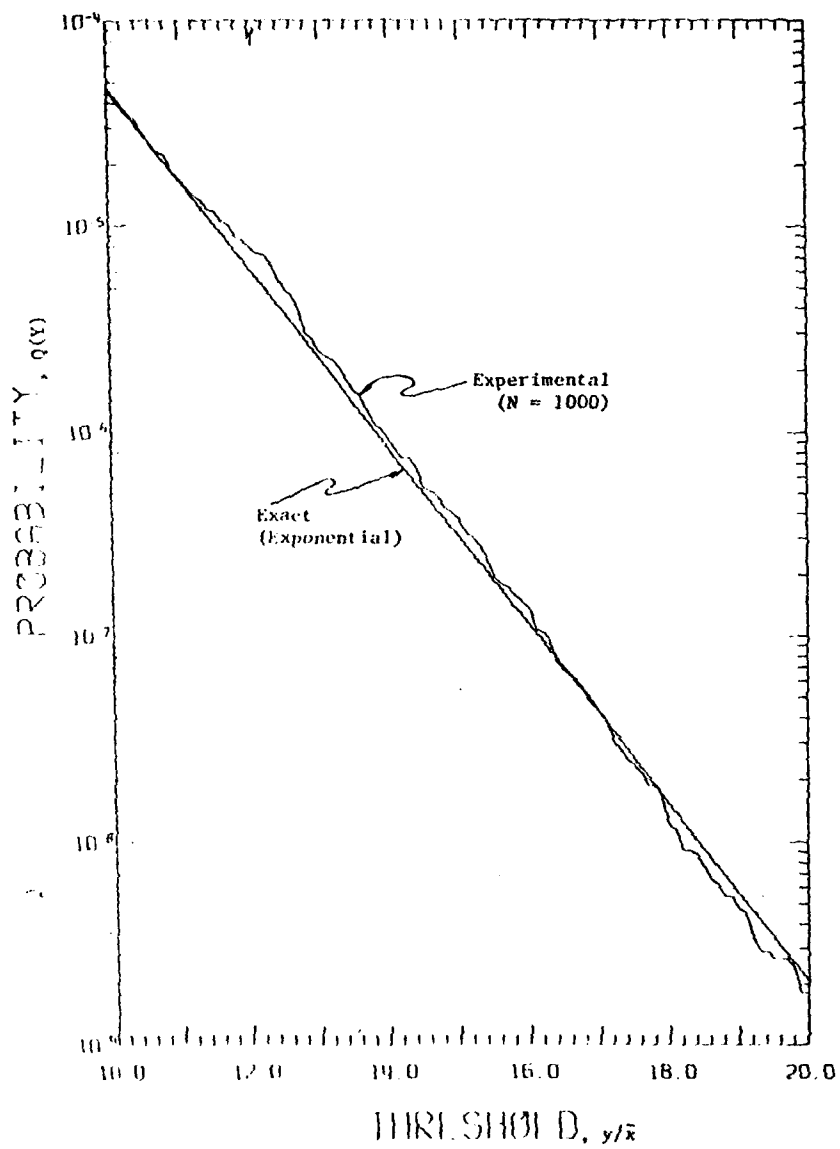


Figure 3. Application of Importance Sampling to Exponential Distribution

and the exact distribution function.* The value of $\bar{x}_m = 13.8\bar{x}$ was used to first compute the histogram, which was then integrated to obtain the sample distribution function. While the procedure was optimized for $Z(Y) \approx 10^{-6}$, we observe that relatively small errors exist throughout the range in probability that covers five orders of magnitude. Thus the procedure is relatively insensitive to the precise value of \bar{x}_m (which could also be established by examining (25)).

Example 2: Gaussian Distribution

Let x be a zero-mean Gaussian random variable, again with $y = F(x) = x$, where

$$p(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-x^2/2\sigma^2}, \quad (28)$$

and

$$Q(Y) = \int_Y^{\infty} p(x) dx. \quad (29)$$

For importance sampling we will modify $p(x)$ by changing (increasing) σ so that

$$p_m(x) = \frac{1}{\sqrt{2\pi}\sigma_m} e^{-x^2/2\sigma_m^2} \quad (30)$$

The weight is given by

$$w(x) = \frac{p(x)}{p_m(x)} = \frac{\sigma_m}{\sigma} e^{-(1/\sigma^2 - 1/\sigma_m^2)x^2/2}. \quad (31)$$

*The inputs to the process were exponentially distributed random variables. If u is a uniformly distributed random variable $(0,1)$, then $x = -\ln(u)$ generates an exponential random variable with unit mean. With importance sampling the precise statistical properties of the pseudo-random numbers, u , are no longer critical as they would be with conventional sampling techniques.

From (21) we obtain for a single observation

$$\bar{z}_m^2 = \frac{\sigma_m/\sigma}{\sqrt{2 - \sigma^2/\sigma_m^2}} Q\left(\frac{Y\sqrt{2 - \sigma^2/\sigma_m^2}}{Q(Y)}\right), \quad (32)$$

where $Q(Y)$ is given by (29). The relative variance for N independent observation is given by

$$\frac{\text{var}[\bar{z}_m]}{(\bar{z}_m)^2} = \frac{1}{N} \left[\frac{\frac{\sigma_m/\sigma}{\sqrt{2 - \sigma^2/\sigma_m^2}} \frac{Q\left(\frac{Y\sqrt{2 - \sigma^2/\sigma_m^2}}{Q(Y)}\right)}{Q^2(Y)} - 1 \right]. \quad (33)$$

In anticipation of the result $\sigma_m \gg \sigma$ we might be tempted to simplify (33) by setting $\sqrt{2 - \sigma^2/\sigma_m^2} = \sqrt{2}$; however, the resulting approximation would not yield an optimum solution as a function of σ_m . As it stands, (33) is best handled numerically.

We can find an optimum value of σ_m that minimizes (33) for a specific value of Y . For example, if $Y = 4.7\sigma$, $Q(Y) \approx 10^{-6}$. A value of $\sigma_m = 4.8\sigma$ was found to be the optimum for this value of Y , although (33) was nearly flat over the interval $4.0 \leq \sigma_m/\sigma \leq 5.5$ so that we can conclude that $\sigma_m = Y$ is essentially the optimum. For $\sigma_m = Y = 4.7\sigma$, the relative variance is

$$\frac{\text{var}[\bar{z}_m]}{(\bar{z}_m)^2} = 49/N, \quad (34)$$

which means that $N = 1000$ observations would result in a relative standard deviation of 22% in the estimate of $Q(Y)$. In Figure 4 we show the comparison between the experimental ($N = 1000$ and $\sigma_m = 4.7\sigma$) and exact distribution functions for the Gaussian case. While σ_m was chosen to optimize the procedure for $Q(Y) \approx 10^{-6}$, we note that the error is relatively small over most of the range in probability plotted in the figure. We see also that the error is somewhat larger than in Figure 3 for the exponential distribution, as we could predict from (21) and (34).

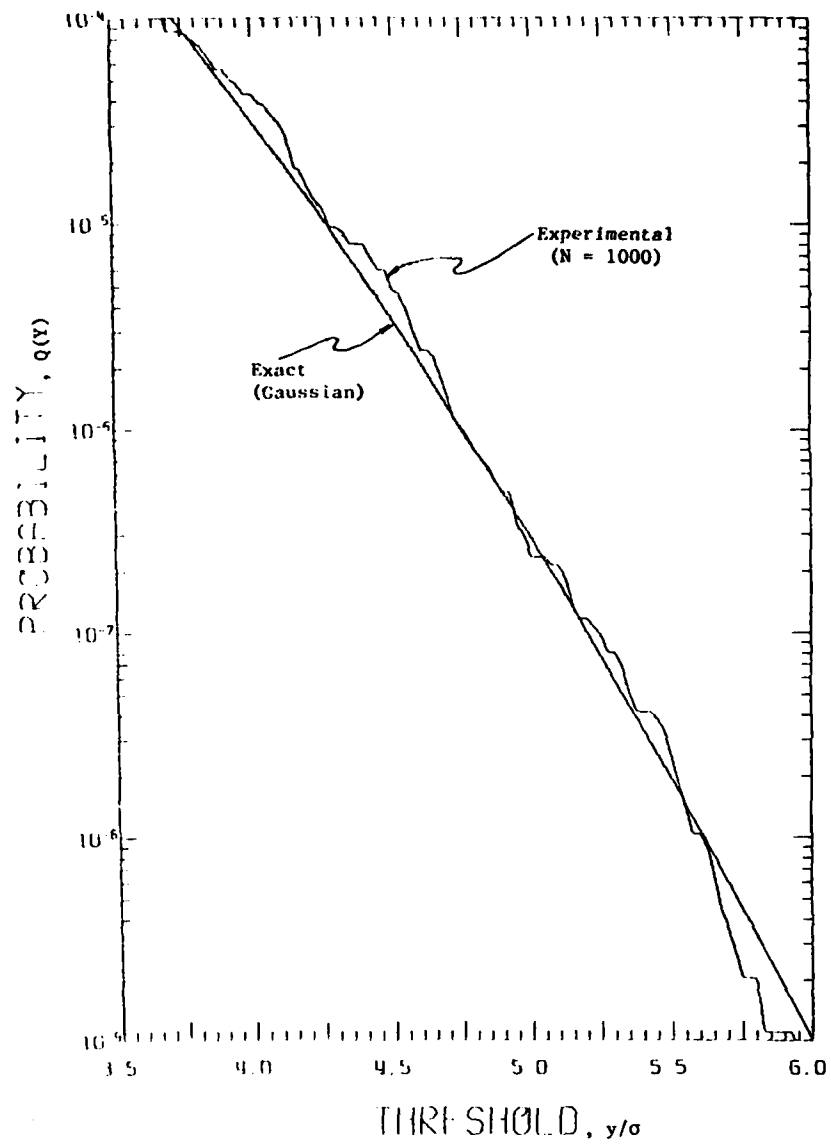


Figure 4. Application of Importance Sampling to Gaussian Distribution

Practically every situation of interest in the analysis of radar and communication systems can be handled by either the exponential or Gaussian random variable as the input to a process (no counter examples are known to the author). For example, the log-normal distribution is a simple transformation of the Gaussian, and the Weibull of an exponential. In fact, one can create an exponential random variable from the Gaussian and vice versa as we will show in the next two examples.

Example 3: Sum-Square of Two Gaussian Variates

Let x_1 and x_2 be two independent, zero-mean Gaussian random variables at the input to our processor, and let

$$y = F(x_1, x_2) = x_1^2 + x_2^2 \quad (35)$$

If we select $\sigma^2 = 0.5$ then y will be an exponential random variable with $y = 1$. With importance sampling the weight applied to each outcome will be a function of only x_1 and x_2 . From (16), (28), and (30) we can write

$$w(x_1, x_2) = 2\sigma_m^2 e^{-(1 - 1/2\sigma_m^2)(x_1^2 + x_2^2)} \quad (36)$$

and with (35)

$$w(x_1, x_2) = 2\sigma_m^2 e^{-(1 - 1/2\sigma_m^2)y} \quad (37)$$

We have applied the above procedure for generating the exponential random variable and with importance sampling we have computed the experimental distribution function. We chose $\sigma_m = 2.63$ (or $2\sigma_m^2 = 13.8$) and $N = 1000$. The experimental result is essentially identical to that in Figure 3, at least in the statistical sense.

Example 4: Generation of Gaussian Variate from Exponential

Let x be an exponential random variable at the input to our processor. Within the processor we will create

$$y = F(x) = \sqrt{2x} \cos \theta \quad (38)$$

where θ is a uniformly distributed random variable $(0, 2\pi)$.^{*} The random variable y will be Gaussian distributed, and if $\bar{x} = 1$ the mean value of y will be zero and the variance unity. With importance sampling the weight applied to each outcome will be a function of only x in this example, since the generation of θ is performed within the processor and will not be considered as an input variable. The weight is given by (20) for $\bar{x} = 1$.

We have generated Gaussian random samples by means of (38) and have used importance sampling to estimate the distribution function in the range of $10^{-4} \leq Q(Y) \leq 10^{-9}$. The result is statistically consistent with that of the direct method in Figure 4. A value of $\bar{x}_m = 22.1$ ($= 4.7^2$) corresponds to optimizing the procedure at $Q(Y) = 10^{-6}$.

With this example we no longer have a unique mapping of x into y , which ought to suggest other possibilities in simulation. In addition, we can configure the processors so that all inputs are either exponential or Gaussian random variables as we will show in Example 9.

Example 5: Sum of Exponential Variates

The sum of K independent random variables that are exponentially distributed is a chi-square random variable with $2K$ degrees of freedom. We can write

$$y = F(x_1, \dots, x_K) = \sum_{k=1}^K x_k, \quad (39)$$

where $\{x_k\}$ is the set of K exponential random variables on input. With importance sampling the weight will be a function of the specific values of x_k . From (16), (17), and (19) we can write

$$\begin{aligned} w(x_1, \dots, x_K) &= (\bar{x}_m/\bar{x})^K \exp \left[-(1/\bar{x} - 1/\bar{x}_m) \sum_{k=1}^K x_k \right] \\ &= (\bar{x}_m/\bar{x})^K e^{-(1/\bar{x} - 1/\bar{x}_m) y} \end{aligned} \quad (40)$$

^{*}The use of (38) with $\sin \theta$ substituted for $\cos \theta$ will produce a second, independent Gaussian random variable.

In Figure 5 we show the experimental distribution obtained by using importance sampling on (39) for $K = 5$ and $N = 1000$ replications. The mean value of the input distribution was modified as $\bar{x}_m = 4.7\bar{x}$, which optimized the procedure at $Q(Y) = 10^{-6}$. Note that the error is extremely small throughout the five orders of magnitude in probability shown in Figure 5.

Example 6: Sum of Exponential Variates with Arbitrary Weights

For the previous example all input random variables were uniformly weighted in (39). Let us generalize the situation by assuming arbitrary weights in the summation as

$$y = F(x_1, \dots, x_K) = \sum_{k=1}^K a_k x_k, \quad (41)$$

where $\{x_k\}$ is the set of independent exponential random variables on input with a mean value \bar{x} . With importance sampling the weight is a function of only the input, so the weight remains the same as (40). The distribution function of y is given by

$$Q(y) = \sum_{k=1}^K \prod_{i \neq k} \frac{a_k^{K-1}}{(a_k - a_i)} e^{-y/a_k \bar{x}}. \quad (42)$$

In Figure 6 we have used importance sampling to estimate $Q(Y)$ for $K=2$ and $N = 1000$. The ratio of weights used in (41) varied from .48/.52 (resulting in essentially a chi-square distribution with 4 degrees of freedom) to .05/.95, with the sum of weights being unity in all cases. For each of the cases, the procedure was optimized in the neighborhood of $Q(Y) \approx 10^{-6}$ (e.g., $\bar{x}_m = 10.7\bar{x}$ for the .25/.75 case).

With this example we can create the following interesting situation: let $K=2$ so that the weight used in importance sampling in (40) will be based on two input variables; however, let one of the weights a_k in (41) be zero so that the output y will be exponentially distributed. From one replication to the other we can even interchange the a_k without affecting the distribution function of y . Does importance sampling work in this case? Indeed it does,

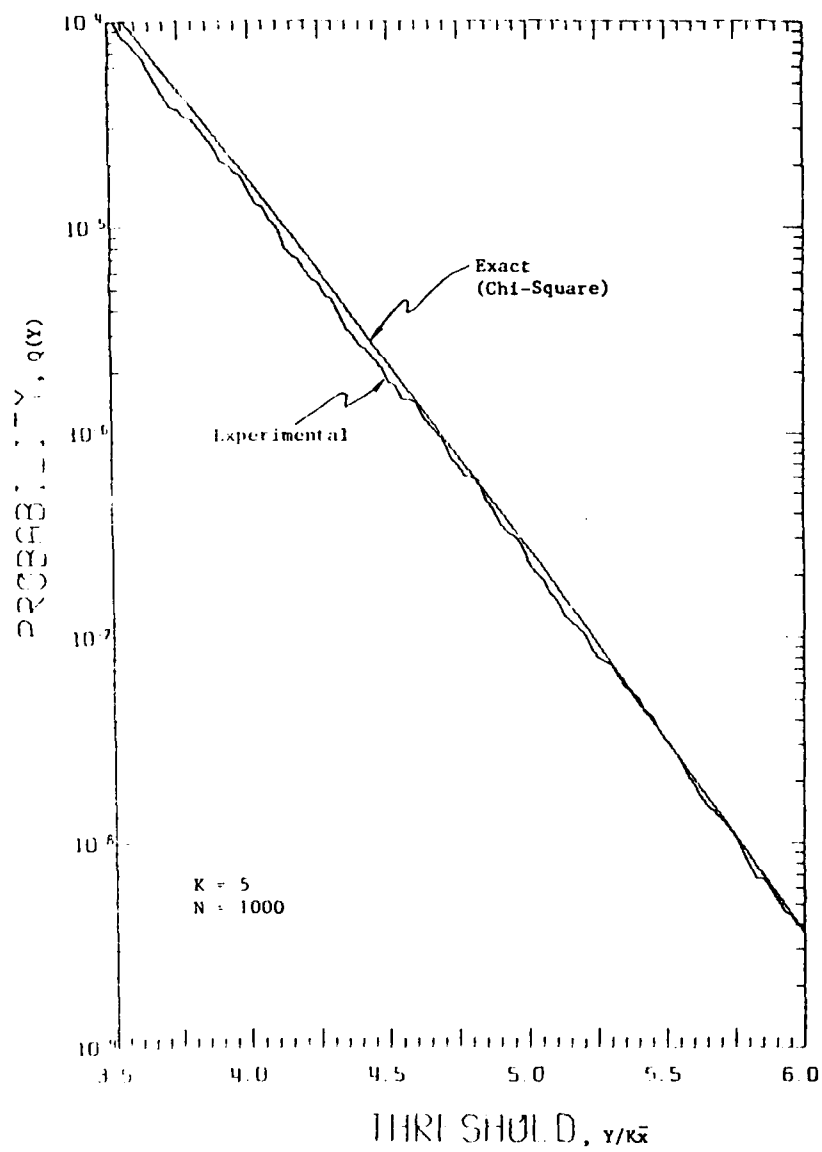


Figure 5. Application of Importance Sampling to Sum of Exponential Variates

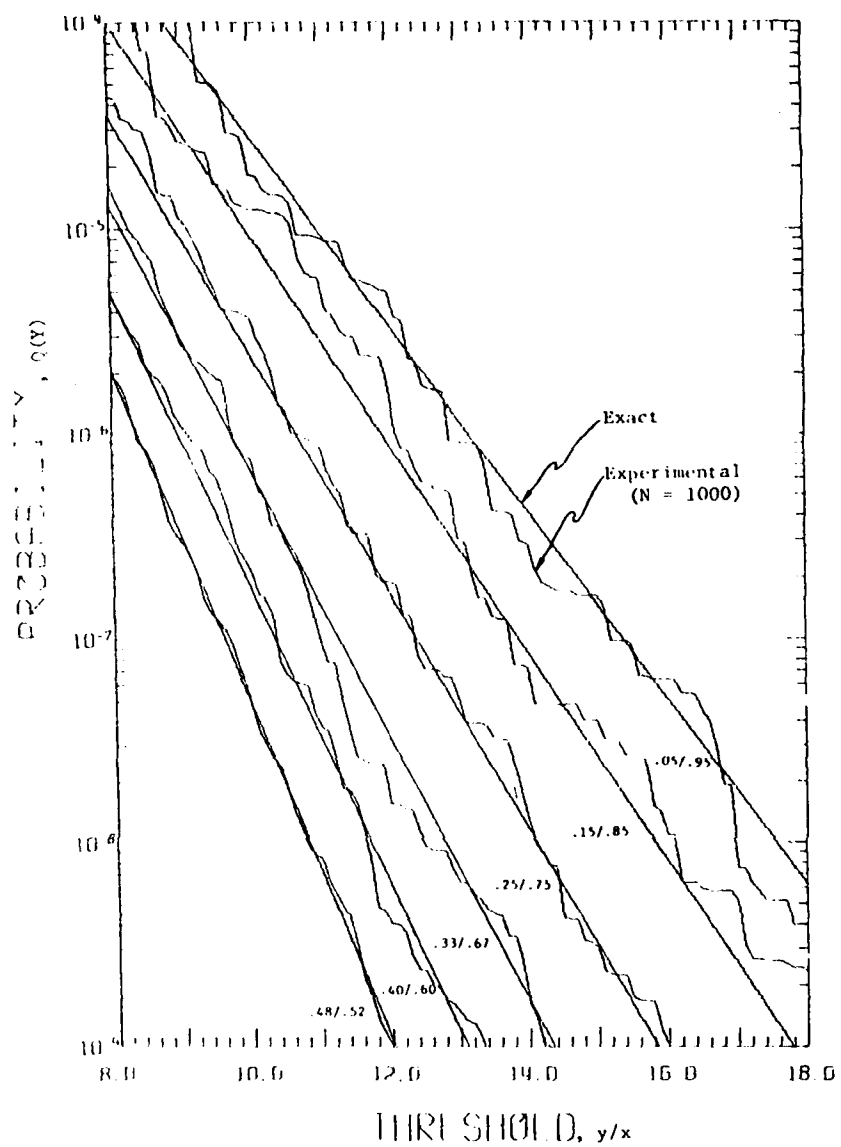


Figure 6. Application of Importance Sampling to Sum of Arbitrarily Weighted Exponential Variates ($K=2$; ratio of weights is indicated)

as we show in Figure 7, with $N = 1000$ and $\bar{x}_n = 13.8$. The error is not as small as it was in Figure 3, but this could be compensated by increasing N .

Example 7: Sum of Log-Normal Variates

The distribution of the sum of log-normal random variables is difficult to compute by conventional numerical methods. Let us begin with an exponentially distributed random variable x_k as one of K inputs to the processor in Figure 1, with $\bar{x} = 1$. We will then generate a Gaussian random variable of unit variance as in Example 4,

$$g_k = \sqrt{2x_k} \cos \theta, \quad (43)$$

where θ is a uniformly distributed random variable $(0, 2\pi)$. The log-normal random variable is generated as

$$l_k = e^{o_L g_k}, \quad (44)$$

where o_L is the standard deviation of the log-variate. The median value of l_k will be unity. The output of the processor will then be

$$y = \sum_{k=1}^K l_k, \quad (45)$$

for which we will use importance sampling to estimate its distribution function. The weight will be given by (40).

Because of the complicated processor above, there is no straightforward way to choose \bar{x}_m which modifies the input distribution for importance sampling. Therefore, we will try some arbitrary values. In Figure 8 we show the experimental distribution functions for eight cases where we have varied x_m from 2 through 30 in steps of 4. In each case $K = 2$, $o_L = 1.0$, and $N = 1000$ observations. We see that all but one of the cases tend to be clustered about a straight line in the format plotted. The exception is the $\bar{x}_m = 2$ case, which

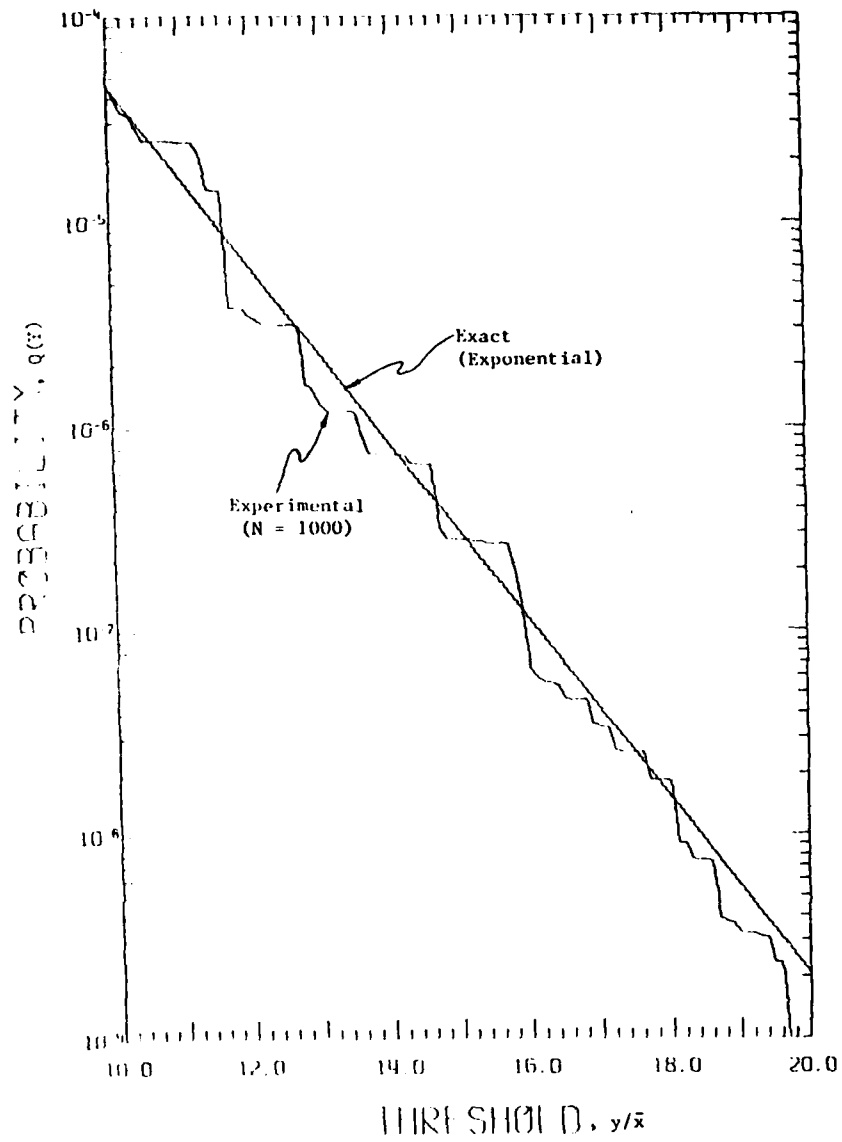


Figure 7. Application of Importance Sampling to Processor Consisting of Two Exponential Variables on Input but only One Used on Output

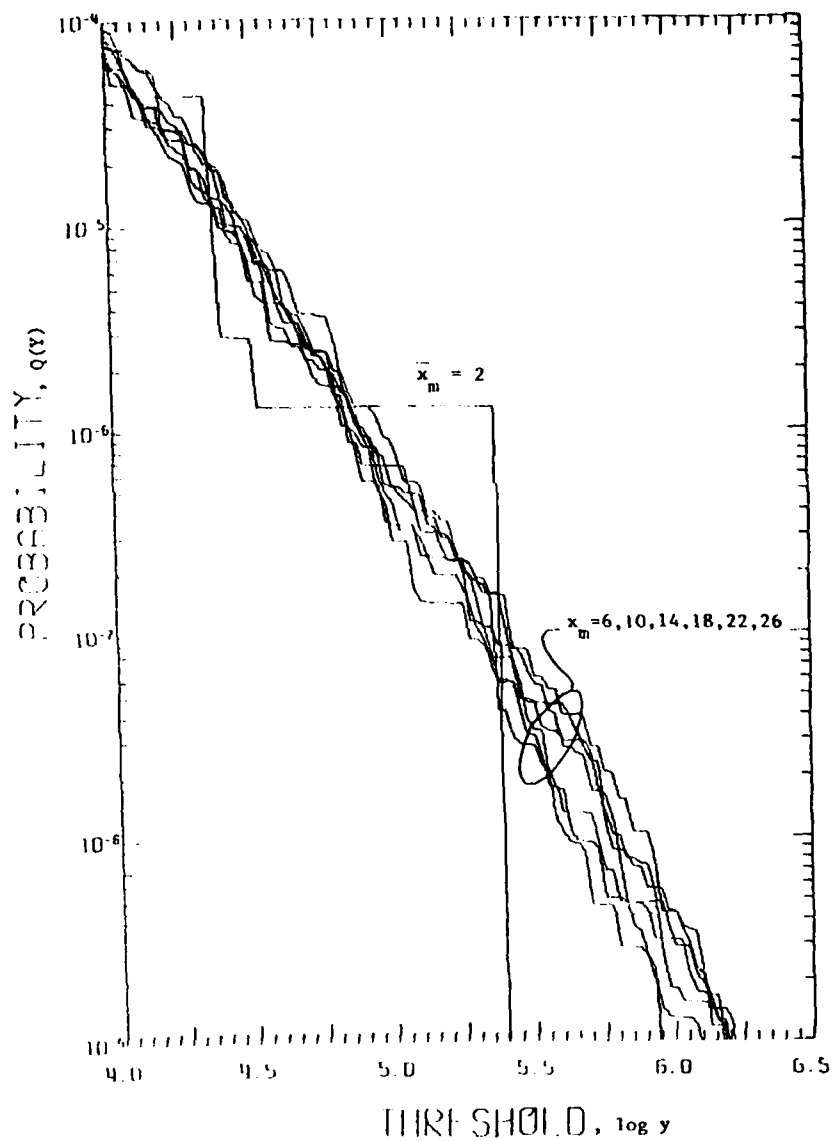


Figure 8. Application of Importance Sampling to Sum of Log Normal Variates ($K = 2$, $\sigma_1 = 1.0$, $N = 1000$)

is too small to produce values of y frequently enough in the range of interest. The remaining cases establish the validity of importance sampling for this example since each estimate is unbiased; the spread of the cluster is a measure of the standard deviation of the estimate. By inference, $\bar{x}_m = 1$ (which is conventional sampling) with a sufficient number of observations would also produce a result that would fall within the cluster in Figure 8. Almost any value of \bar{x}_m , especially in the interval $6 \leq \bar{x}_m \leq 30$, could be used with importance sampling to produce an acceptable result. However, since $\bar{x}_m = 22$ optimizes the procedure for a single Gaussian random variable at $Q(Y) = 10^{-6}$, there is probably no requirement to exceed this value in the general case. We could also repeat the above procedure whenever K or σ_L were changed.

We have also used importance sampling on (45) by beginning with independent Gaussian random variables as the input to the processor. The above conclusions were unaffected by this change, however (the range $2 \leq \sigma_m \leq 8$ proved satisfactory).

Example 8: Sum of Log-Normal Phasors

A situation that is common in the analysis of radar clutter is the summation of random phasors where the amplitude of each is log-normally distributed. We will generate the log-normal random amplitude ℓ_k as in the previous section (but σ_L now refers to the signal amplitude) and then generate the random phasor as

$$v_k = \ell_k e^{j\phi_k} \quad (46)$$

where ϕ_k is a uniformly distributed random variable $(0, 2\pi)$. Finally, we form

$$y = \left| \sum_{k=1}^K v_k \right|^2 \quad (47)$$

for which we will use importance sampling to estimate its distribution function.

In Figure 9 we show the experimental distribution functions for three cases, $\bar{x}_m = 18, 22$, and 26 , where $K=2$, $\sigma_L = 0.5$, and $N = 1000$ for all cases. The experimental curves again cluster fairly closely throughout the entire range in probability plotted.

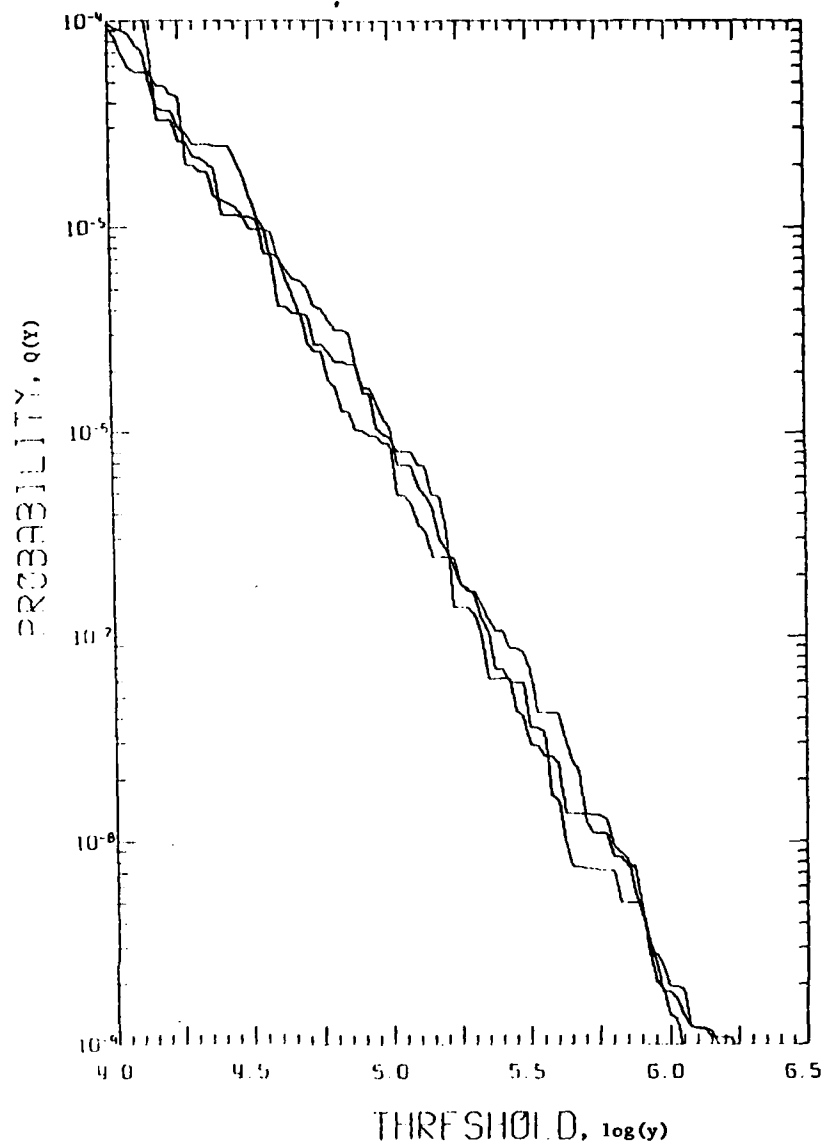


Figure 9. Application of Importance Sampling to Sum of Log-Normal Phasors ($K = 2$, $\sigma_L = 0.5$, $N = 1000$)

Example 9: Mixed Statistics

In Examples 5 through 8 we have superimposed random samples from the same statistical distribution. But in radar and communication systems, signals are often combined with mixed statistics. For example, radar clutter, which is often assumed to be log-normally distributed, would be combined with thermal noise, which is Rayleigh-amplitude distributed. In this example we will simulate this case and use importance sampling to estimate the distribution function. We will again begin with two exponentially distributed random variables as the input to the processor in Figure 1. A log-normal phasor, V_1 , will be generated from one of the exponentially distributed inputs, x_1 , by means of (43), (44), and (46). The Rayleigh-amplitude phasor will be generated from the second exponentially distributed input, x_2 , as

$$V_2 = \sqrt{x_2} (\cos \phi + j \sin \phi) \quad , \quad (48)$$

where ϕ is a uniformly distributed random variable $(0, 2\pi)$ that is generated within the processor. The output of the processor will be

$$y = |V_1 + aV_2|^2 \quad , \quad (49)$$

where a is a factor used to scale one process with the other. We will constrain $\bar{x} = 1$ for both inputs.

In Figure 10 we show the application of importance sampling used to estimate the distribution function of y for $a = 3$ and $N = 10^4$ observations. The four values of $\bar{x}_m = 14, 18, 22$, and 26 were used to generate the four experimental curves. Note that they are tightly clustered, especially for the higher probabilities shown in the figure. We have also shown the exact distribution functions for the components that make up the sum in (49).

Since the four experimental curves in Figure 10 are so closely clustered throughout five orders of magnitude in probability, even with such widely different values of the importance sampling scaling parameter, \bar{x}_m , we can conclude with high probability that the true distribution also lies within the cluster. Thus we can also conclude that importance sampling works when the process

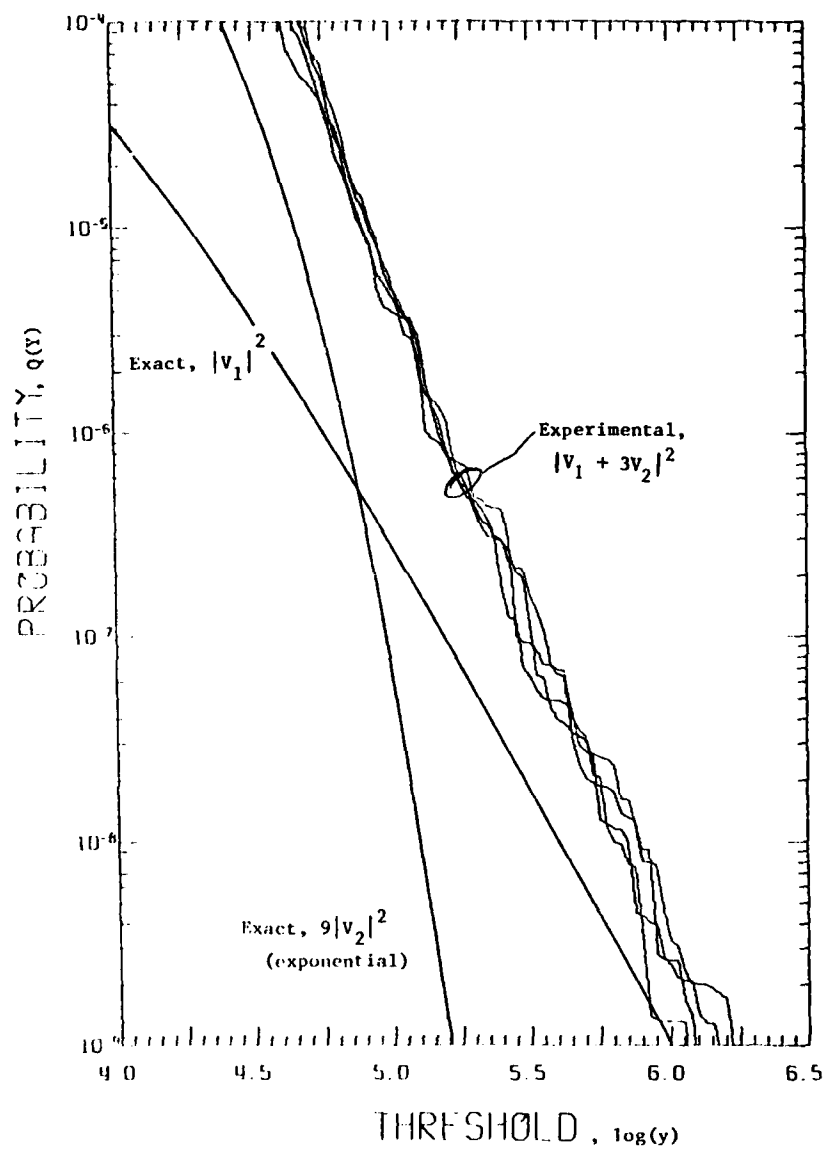


Figure 10. Application of Importance Sampling to Example of Mixed Statistics ($N = 10^4$)

under investigation involves a combination of random variables with different statistics. In the application of importance sampling one probably should create all random variables by some transformation of a set of input random variables with a common distribution.

Example 10: CFAR

A simple approach in radar to achieving a constant false alarm rate (CFAR) in the presence of nonstationary noise is to set the detection threshold on the basis of the average noise power in a number of reference samples where each of these samples is assumed to represent noise only. Such a scheme is denoted as cell averaging CFAR. The output of the processor might be

$$y = \frac{x}{\sum_{i=1}^R x_i} \quad (50)$$

where x is the sample under test and the set of samples $\{x_i\}$ form the CFAR reference. Note that if we modify all $R+1$ input variables equally by changing the common mean value, the output y remains unchanged and importance sampling would seem to be inapplicable in this case. One could apply importance sampling to y if $p(y)$ were known; however, this would be a severe limit to the approach.

In order to find a method that works on the input variables it was decided to treat the numerator in (50) as the input process; the denominator will be generated internally to the processor, just as we generated θ in Example 4. The reasoning behind this choice is that the variance of the denominator is much less than the numerator, and the low probability behavior of the numerator has a much greater impact on the outcome than does the denominator. However, unless R is reasonably large we will not get much of an improvement with importance sampling compared to the conventional approach. If the random variables are exponentially distributed, the following number of observations N will be needed for each value of R to provide the same quality of result with importance sampling [$Q(y) = 10^{-6}$]:

R	∞	200	100	50	20
N	1000	2300	4700	15,000	170,000

In most applications of cell averaging CFAR in radar we will also employ some noncoherent integration of samples prior to forming the CFAR ratio. In such cases we can write the output as

$$y = \frac{\sum_{k=1}^K x_k}{\sum_{i=1}^R x_i} \quad (51)$$

Usually there will be M reference samples for each sample in the numerator, so we can write $R = KM$. In a typical situation we might have $K = 6$ and $M = 10$ so that $R = 60$, which permits us to use a reasonably small number of replications to determine $Q(Y)$ with importance sampling applied to only the numerator of (51). In Figure 11 we show the results of this case with $N = 10,000$ observations; all random variables are exponentially distributed and $\bar{x}_m = 5\bar{x}$ for the $K = 6$ input variables. Thus importance sampling also works with cell averaging CFAR, but we have been forced to redefine the input to the processor.

2.4 CONCLUSIONS

Importance sampling has wide application in the simulation of signals in radar and communication systems. It is robust and efficient, producing reliable estimates of the low-probability tail of the distribution function with typically 1000 replications of the experiment. It has been shown to work with multiple input processes when all inputs belong to the same distribution and are distorted equally. For a process involving a combination of several types of signals, each from a different statistical family, one can still apply importance sampling by redefining the procedure in which the signals are generated. In effect, a new processor is created so that all input signals will

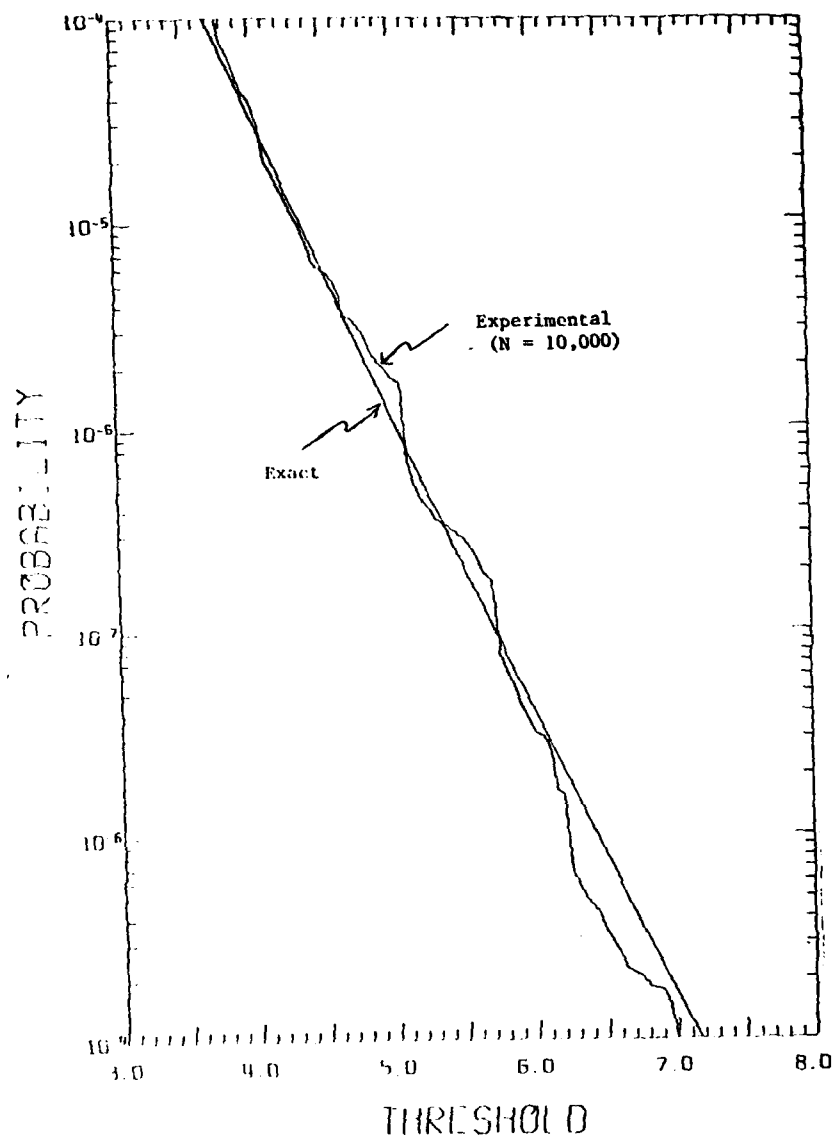


Figure 11. Application of Importance Sampling
to CFAR Example

belong to the same statistical family and all will be distorted equally in importance sampling. In some cases, such as with cell averaging CFAR, we will even have to redefine the input to the processor. It has been shown that one can begin with either Gaussian or exponentially distributed signals and generate practically any combination of statistical signals used in the simulation of radar and communication systems. We are unaware of any situation of practical interest where this procedure would fail.

3. MODULATION OF CLUTTER WITH A SCANNING ANTENNA

There are two basic ways to scan the antenna in a ground based search radar: either continuously or step-wise. In the latter case we can assume that there is no antenna motion relative to the ground for a fixed time which we designate as the on-target time. Received pulses from the ground clutter will be partially correlated which is due entirely to the internal motion of the clutter itself (the wind-induced motion of trees, etc.). If the antenna moves from pulse to pulse, then the return from any point on the ground will be amplitude modulated as a result of the time-varying antenna gain in the direction of that point. The correlation properties of ground clutter will then be affected by the scanning modulation in addition to the internal motion of the clutter. In many cases, the scan modulation effect dominates the internal motion effect, and for simulating clutter signals the scan modulation severely complicates an otherwise straightforward implementation.

For the purposes of this discussion we will distinguish between two types of radar signal processing: continuous and batch. With continuous processing the radar generally utilizes all past data in a recursive filter to process each pulse, and there is usually one output signal for each pulse. Continuous processing is often implemented in continuously scanning systems, at least for older radars. With batch processing the radar collects a sequence of pulses before any processing is implemented, and the processing from one batch to the next is independent. Batch processing is implemented on all step scan systems (as known to the author) and some of the newer continuous scan systems, especially the 3-D search radars utilizing an electronic scan in elevation. The FFT processor is one example of a batch processor.

The distinction between continuous and batch processing has a dramatic effect on the techniques used to simulated clutter signals, so much so that the techniques and computer algorithms used for one system will bear little similarity to those used for the other. In effect, there will be two distinctly different computer programs required to simulate both cases. The reason is that the continuous processing

system demands almost a pulse-by-pulse approach to computing the interaction of the antenna with the clutter geometry, while the batch processing system usually needs to be updated only once for each batch of pulses. The latter case is simpler and cheaper to simulate, and few questionable approximations will be required out of necessity. For this reason we will concentrate our discussion on the continuous processing system. We will assume in the following analysis that the radar has 2-D resolution and the antenna has a fan beam in elevation.

3.1 FORMULATION OF THE CLUTTER SIGNAL

Let us designate the complex signal associated with the k^{th} scatterer in a range resolution cell as $V_k(t)$, which we have made time-varying in order to account for internal motion of the clutter. In general, $|V_k(t)|^2$ will have the dimensions of radar cross section (RCS), or some power scaled version of it according to the radar range equation. In addition, let $g(\theta)$ be the one-way voltage gain of the antenna in the azimuth, θ , direction. The composite signal (complex voltage) received by the radar when the antenna scans continuously in azimuth will be proportional to

$$V(t) = \sum_k V_k(t) g^2(\theta_0 + \dot{\theta}t - \theta_k) \quad (52)$$

where θ_0 is some reference angle, $\dot{\theta}$ is the scan rate, θ_k is the azimuth of the k^{th} scatterer, and the summation is performed over all scatterers within a range resolution cell. Differences in range among scatterers that would effect the phase of superposition in (52) are implicitly included in the definition of $V_k(t)$.

As written (52) is nontrivial to implement because $V_k(t)$ is a two dimensional process: spatial and temporal. Without any simplification (which can be done only through approximations) we must evaluate (52) for each pulse, and at the very least we must have one scatterer (in a range ring) for each pulse throughout a 360° scan. For a PRF of 400 Hz and a 10 sec scan we will have 4000 total pulses and at least 4000 scatterers. The process $V_k(t)$ is thus described by a

4000 \times 4000 matrix (neglecting the "edge effects" at the beginning and end of the scan). Clearly, some kind of approximation is in order at this point to reduce the size of the problem.

One approach would be to assume that $g(\theta)$ is negligibly small for $|\theta|$ larger than some value (where $\theta = 0$ corresponds to the mainlobe axis.) We could simulate only the mainlobe or the near sidelobes of the antenna, and assume that there is no contribution from sidelobes further out than some point. The problem is that in some cases (in older systems) just simulating the mainlobe might be adequate but in other cases (in newer, high-performance systems) we might have to include many sidelobes. Just where we should stop simulating sidelobes is not readily apparent since it depends on the system being simulated. Few ground rules can be given either because very little analysis has been performed on this subject.

Another approach to simplifying (52) is to focus our attention on $V_k(t)$. In general, this process can be assumed to be random (for the objective of determining detection performance), with the statistics of one scatterer being independent from any other. The random process will be spatially uncorrelated. It is fairly common in the literature to break up the scattering properties of ground clutter into two components: an ac component and a dc component. The ac component, the time-varying one, is the result of the motion of trees, brush, grass, etc., and is generally made up of many individual scatterers (e.g., the leaves on a tree) where no single one dominates. Such a process is easily justified to be Gaussian, at least within a relatively small area on the ground. On the other hand, the dc component is associated with rigid objects, such as bare ground, tree trunks, man-made structures, etc., and non-Gaussian statistics usually prevail in this case. In general, the dc component is stronger (higher power) than the ac component.

With the above distinction between the two components let us define

$$V_k(t) \approx V_k + \tilde{V}_k(t) \quad (53)$$

where V_k is the dc component of the k^{th} scatterer and $\tilde{V}_k(t)$ is the ac component. If we substitute (7) into (1) we obtain

$$V(t) = \sum_k V_k g^2(\theta_0 + \dot{\theta}t - \theta_k) + \sum_k \tilde{V}_k(t) g^2(\theta_0 + \dot{\theta}t - \theta_k) \quad (54)$$

At first glance it appears that we have complicated the situation because we now have two summations instead of one. However, the first one is straightforward; it is a circular convolution, something that is easily implemented. For the second summation, which has all the inherent complexity of (52) we will truncate the antenna pattern as described previously, but now the effect of the truncation will be much less significant since $\tilde{V}_k(t)$ will usually be much weaker than $V_k(t)$.

Equation (54), or at least the second term of it, is not particularly economical to implement in terms of computer time, especially when computing false-alarm performance. For this reason we will pursue further approximations.

3.2 FURTHER APPROXIMATIONS

In general (52) and (54) describe a nonstationary process because clutter samples in one azimuth sector can be much stronger than in another. This is another way of saying that clutter is spatially nonhomogeneous. If this property of clutter applies to both terms in (54) then there is little more we can do to simplify the expression. However, if the ac component of clutter is homogeneous (mere speculation at this time) we can treat the second term in (54) as a stationary process. With this assumption we can apply Fourier analysis to determine the spectrum of this term which we rewrite as

$$V_{ac}(t) = \sum_k \tilde{V}_k(t) g^2(\theta_0 + \dot{\theta}t - \theta_k) \quad (55)$$

In determining the second moment, let us write

$$V_{ac}(t_1) V_{ac}^*(t_2) = \sum_k \tilde{V}_k(t_1) \tilde{V}_k^*(t_2) g^2(\theta_0 + \dot{\theta}t_1 - \theta_k) g^2(\theta_0 + \dot{\theta}t_2 - \theta_k) \quad (56)$$

When we take the ensemble average of (56) the cross terms vanish as

$$\tilde{V}_k(t_1) \tilde{V}_l^*(t_2) = 0, \quad k \neq l$$

so that

$$\underbrace{v_{ac}(t_1)v_{ac}^*(t_2)}_{\text{autocorrelation function of composite process}} = \sum_k \underbrace{\tilde{v}_k(t_1)\tilde{v}_k^*(t_2)}_{\text{autocorrelation function of clutter internal motion}} g^2(\theta_0 + \dot{\theta}t_1 - \theta_k) g^2(\theta_0 + \dot{\theta}t_2 - \theta_k) \quad (57)$$

The autocorrelation functions above will be assumed to be nonstationary (even though we will later assume that the autocorrelation function of the clutter internal motion is stationary). Thus we will define

$$\tilde{R}(t_1, t_2) = \tilde{v}_k(t_1)\tilde{v}_k^*(t_2) \quad (58)$$

$$R_{ac}(t_1, t_2) = v_{ac}(t_1)v_{ac}^*(t_2) \quad (59)$$

and

$$h_k(t) = g^2(\theta_0 + \dot{\theta}t - \theta_k) \quad (60)$$

Note that we are assuming that (58) is independent of the subscript k so that the clutter process is spatially homogeneous. With this notation, (57) becomes

$$R_{ac}(t_1, t_2) = \sum_k \tilde{R}(t_1, t_2) h_k(t_1) h_k(t_2) \quad (61)$$

At this point we will introduce the two-dimensional Fourier transform as used in Reference 7, so that

$$\tilde{\Gamma}(f_1, f_2) = \iint \tilde{R}(t_1, t_2) e^{-j2\pi(f_1 t_1 - f_2 t_2)} dt_1 dt_2, \quad (62)$$

and

$$\tilde{R}(t_1, t_2) = \iint \Gamma(f_1, f_2) e^{j2\pi(f_1 t_1 - f_2 t_2)} df_1 df_2, \quad (63)$$

with a similar set of expressions for $\Gamma_{ac}(f_1, f_2) \leftrightarrow R_{ac}(t_1, t_2)$. The two-dimensional Fourier transform of (61) is now

$$\Gamma_{ac}(f_1, f_2) = \sum_k \iint \tilde{R}(t_1, t_2) h_k(t_1) h_k(t_2) e^{-j2\pi(f_1 t_1 - f_2 t_2)} dt_1 dt_2. \quad (64)$$

Let $t_2 = t_1 - \tau$ and assume that $\tilde{R}(t_1, t_2) = \tilde{R}(\tau)$, which is the definition of a stationary process. After some manipulation we can write

$$\Gamma_{ac}(f_1, f_2) = \sum_k \int \tilde{R}(\tau) e^{-j2\pi f_2 \tau} \left\{ \int h_k(t_1) h_k(t_1 - \tau) e^{-j2\pi t_1 (f_1 - f_2)} dt_1 \right\} d\tau. \quad (65)$$

The quantity within the braces is the ambiguity function of the waveform $h_k(t)$. Let us define

$$\chi_k(\tau, f) = \int h_k(t) h_k(t - \tau) e^{-j2\pi f t} dt, \quad (66)$$

so that (65) becomes

$$\Gamma_{ac}(f_1, f_2) = \sum_k \int \tilde{R}(\tau) \chi_k(\tau, f_1 - f_2) e^{-j2\pi f_2 \tau} d\tau. \quad (67)$$

This is about as far as we can conveniently go without solving for a specific example.

3.3 THE GAUSSIAN BEAM

If we assume that the antenna beamshape is Gaussian then the integrations become relatively tractable. In particular, let

$$h_k(t) = e^{-at^2} \quad (68)$$

for the moment. Then it can be shown that

$$X_k(\tau, f) = \sqrt{\frac{\pi}{2a}} e^{-\pi^2 f^2 / 2a} e^{-a\tau^2 / 2} e^{-j\pi f \tau} \quad (69)$$

Now if the antenna pattern is defined by

$$g^2(\theta) = e^{-(\alpha\theta/\theta_{3dB})^2} \quad (70)$$

where $\alpha = 1.6651$ and θ_{3dB} is the one-way half-power beamwidth, we note that from (60) and (68)

$$a = (\alpha\dot{\theta}/\theta_{3dB})^2 \quad (71)$$

and

$$\tau \rightarrow \tau + (\theta_o - \theta_k)/\dot{\theta} \quad (72)$$

Thus (69) is given by

$$X_k(\tau, f) = \sqrt{\frac{\pi}{2}} \frac{\theta_{3dB}}{\alpha\dot{\theta}} e^{-\pi^2 f^2 (\pi\theta_{3dB}/\alpha\dot{\theta})^2} e^{-\pi^2 (\alpha\dot{\theta}/\theta_{3dB})^2 \tau^2} e^{-j2\pi f ((\theta_o - \theta_k)/\dot{\theta} - \tau/2)} \quad (73)$$

Note that in (67) the subscript k is only associated with the ambiguity function, and in (73) the subscript k only appears in the last term. Thus if we sum (73) over k we must evaluate the following expression

$$\sum_k e^{-j2\pi f ((\theta_o - \theta_k)/\dot{\theta} - \tau/2)} \quad (74)$$

However, as the increments become small so that the summation approaches an integration, (74) will have a significant value only for $f \approx 0$. As the limits on the integration get larger and larger, (74) becomes more like a

δ -function, $\delta(f)$. For our purposes we will assume that it is a δ -function because we will never employ long enough integration times to be able to resolve the difference in shape of (74) from a δ -function. We must also include a scale factor in the evaluation of (74) as

$$\sum_k e^{j2\pi f((\theta_0 - \theta_k)/\dot{\theta} - \tau/2)} = \frac{\dot{\theta}}{\Delta\theta} \delta(f) \quad (75)$$

where $\Delta\theta$ is the azimuth spacing of the clutter samples around the range ring (or the average spacing if the samples are nonuniformly spaced). With this assumption the summation of (73) over k becomes

$$\sum_k \chi_k(\tau, f) = \sqrt{\frac{\pi}{2}} \frac{\theta_{3dB}}{\alpha\dot{\theta}} e^{-\frac{1}{2}(\alpha\dot{\theta}\tau/\theta_{3dB})^2} \delta(f) \quad (76)$$

Furthermore, (67) reduces to

$$\Gamma_{ac}(f_1, f_2) = \delta(f_1 - f_2) \cdot \int \tilde{R}(\tau) \chi(\tau) e^{-j2\pi f_2 \tau} d\tau \quad (77)$$

where

$$\chi(\tau) = \sqrt{\frac{\pi}{2}} \frac{\theta_{3dB}}{\alpha\dot{\theta}} e^{-\frac{1}{2}(\alpha\dot{\theta}\tau/\theta_{3dB})^2} \quad (78)$$

Note that (77) is just the autocorrelation function of the beam scan modulation function in (60) with the Gaussian beam assumption. Furthermore, (77) is singular along the $f_1 = f_2$ axis, which means that the resultant process is stationary [7]. If we designate $S_{ac}(f)$ as the power spectrum of this stationary process, then we can write

$$S_{ac}(f) = \int \tilde{R}(\tau) \chi(\tau) e^{-j2\pi f \tau} d\tau$$

Now let us define the following Fourier transform pairs

$$S_{im}(f) \leftrightarrow \bar{R}(\tau) \quad (80)$$

$$G(f) \leftrightarrow \chi(\tau) \quad (81)$$

so that the spectrum of the composite process in (79) can be written as

$$S_{ac}(f) = S_{im} \star G(f) \quad (82)$$

where $S_{im}(f)$ is the spectrum of the clutter internal motion, $G(f)$ is the scan modulation spectrum, and the star designates a convolution.

Let us continue with the assumption of a Gaussian beam. We will define

$$f_g = \frac{\alpha^2 \theta_0}{\sqrt{2} \pi \theta_{3dB}} \quad (83)$$

so that when we take the Fourier transform of (77) we obtain

$$G(f) = \sqrt{\frac{\pi}{2}} \frac{\theta_{3dB}}{\alpha \theta_0} \underbrace{\frac{\alpha}{\sqrt{\pi} f_g} e^{-(\alpha f / f_g)^2}}_{\text{normalized to unit power}} \quad (84)$$

normalized to unit power

Note that f_g is the two-sided half-power spectral width. Let us also assume that $R(\tau)$ is a Gaussian-shaped autocorrelation function

$$R(\tau) = \frac{P_T \Lambda \theta_0}{2\pi} e^{-(b\tau)^2} \quad (85)$$

Note that this autocorrelation function applies to a single clutter sample; therefore we have chosen P_T as the average power associated with all clutter within a range ring throughout 2π radians in azimuth. The Fourier transform of (85) is also Gaussian shaped as

$$S_{im}(f) = \frac{P_T \Lambda \theta_0}{2\pi} \underbrace{\frac{\alpha}{\sqrt{\pi} f_{im}} e^{-(\alpha f / f_{im})^2}}_{\text{normalized to unit power}} \quad (86)$$

normalized to unit power

where

$$b = \pi f_{im} / \alpha$$

and f_{im} is the two-sided half-power width of the internal motion spectrum.
The convolution of (84) and (86) is

$$S_{ac}(f) = P_T \frac{\sqrt{\pi}}{\alpha} \frac{0.3dB}{\sqrt{2} \cdot 2\pi} \underbrace{\frac{\alpha}{\sqrt{\pi} f_{ac}} e^{-(\alpha f / f_{ac})^2}}_{\text{normalized to unit power}} \quad (87)$$

\uparrow
 1.064
 total power in range ring

\uparrow
 2-way half-power beamwidth : 2π

where

$$f_{ac}^2 = f_{im}^2 + f_g^2 \quad (88)$$

We can neglect the scan modulation when $f_{im} \gg f_g$, or from (83) when

$$f_{im} > .624 \dot{0}/0_{3dB} \quad (89)$$

We can replace the \gg sign with a > 4.0 without any noticeable effect on performance.

4. GENERATING CLUTTER SEQUENCES FOR A GROUND-BASED RADAR

In Reference 8 several general techniques are described for generating clutter sequences. They are all based on properties of the discrete Fourier transform and the fact that samples in the frequency domain will be statistically independent of each other. We will generate a set of independent random phasors $\{X(n\Delta f)\}$ at uniformly spaced increments, Δf , in the frequency domain, such that

$$\overline{|X(n\Delta f)|^2} = \Delta f S(n\Delta f) \quad (90)$$

where $S(f)$ is the desired power spectral density and the bar on the left side of (90) designates an ensemble average. The amplitude distribution of the phasors $\{X(n\Delta f)\}$ need not be Rayleigh [8]. The correlated time sequence is obtained by taking the discrete Fourier transform as

$$x(k\Delta t) = \sum_n X(n\Delta f) e^{j2\pi kn\Delta f\Delta t} \quad (91)$$

Usually the sample spacing in the time domain, Δt , will be given (e.g., the pulse repetition period in a pulsed radar), but Δf is under our control. Before we define how Δf should be chosen we note that the time sequence repeats with a period given by

$$T_r = 1/\Delta f \quad (92)$$

Let us also define the two-sided half-power width of the spectrum as f_{3dB} . The correlation distance is then approximately

$$T_c = 1/f_{3dB} \quad (93)$$

Since the time sequence repeats with a period T_r , we can utilize only a portion of the period without having the beginning of the desired sequence being correlated with the end. We should therefore choose

$$T_r \geq T + T_c \quad (94)$$

where T is the duration of the desired sequence. If we make use of (92) and (93) we can rewrite (94) as

$$\Delta f \leq \frac{f_{3dB}}{1 + f_{3dB}T} \quad (95)$$

There is one other constraint on Δf : We must be able to resolve the spectrum. From Reference 8 (p. 120) we must have $\Delta f \leq 0.63 f_{3dB}$ for a Gaussian-shaped spectrum. The constant will be slightly different for other spectral shapes, but it will probably not be less than 0.63 for ground-based radar clutter spectra. Therefore we can write

$$\Delta f \leq 0.63 f_{3dB} \cdot \min \left\{ 1, \frac{1.60}{1 + f_{3dB}T} \right\} \quad (96)$$

In general, $f_{3dB}T$ will be less than 0.60 for a ground-based surveillance radar, so the simple constraint $\Delta f \leq 0.63 f_{3dB}$ will apply in most cases; however, (96) will always be applicable.

There are two general ways to implement (91): the fast Fourier transform (FFT) and brute-force approaches. We will now discuss each.

4.1 FFT APPROACH

Let us define the repetition frequency as

$$f_r = 1/\Delta t \quad (97)$$

Next, let us divide this interval into N_r equal increments so that

$$\Delta f = f_r / N_r \quad (98)$$

From (97) we note that

$$N_r = 1/\Delta t \Delta f \quad (99)$$

so that (91) can be written as

$$x(k\Delta t) = \sum_{n=0}^{N_r-1} x(n\Delta f) e^{j2\pi kn/N_r} \quad (100)$$

which is the conventional definition of the discrete Fourier transform that can be implemented as an FFT.

In order to determine how large N_r should be let us utilize (98) and the inequality $\Delta f \leq 0.63 f_{3dB}$, and write

$$N_r \geq 1.60 f_r / f_{3dB} \quad (101)$$

with the understanding that N_r might have to be even larger if (95) were to apply. Now let us work with some numbers. A general rule of thumb is that the spectral width of ground clutter for a ground-based radar with a non-scanning antenna is about 3% of the maximum wind Doppler [8]. Thus we can write

$$f_{3dB} = .06 V_w / \lambda$$

where V_w is the wind velocity and λ is the wavelength. For $V_w = 10$ m/sec and $\lambda = .12$ m (S-band) we have $f_{3dB} = 5$ Hz. If the PRF is 1 kHz ($f_r = 1000$ Hz) then (101) becomes $N_r \geq 320$, which is a large number, especially considering that we might utilize at most only about 16 samples in the time domain. The disadvantage of the FFT approach, at least as defined so far, is that most of the spectral samples will be zero. Even though the FFT is efficient, it must still implement the multiplies by zero. The next approach avoids this shortcoming.

4.2 BRUTE-FORCE APPROACH

The brute-force approach is a direct implementation of (91). We define the samples in the frequency domain only over a limited region of the power spectral density. If we utilize $\Delta f = 0.63 f_{3dB}$, then as few as 5 or 7 samples will be sufficient to define the spectral process (as determined in Reference 8 for Gaussian shaped spectra; the rule might be different for other spectral shapes). Moreover (91) is implemented only for the desired number of time samples. If N_f and N are the number of frequency and time domain samples, respectively, then the computation time will be proportional to the product NN_f , instead of $N_r \log_2 N_r$ for the FFT approach. But since $N \ll N_r$ and $N_f \ll N_r$ will usually be true for ground-based radars, the brute force approach will usually be faster than the FFT approach for generating clutter spectra.

4.3 FFT APPROACH WITH INTERPOLATION

If $f_r \gg f_{3dB}$, as it is with most ground-based radars, then many consecutive time-domain samples will be correlated. We can increase the sample spacing Δt in (91) or (100) to reduce the computation time for the Fourier transform, and then utilize interpolation to obtain samples at the desired rate. Let us define the reduced sample spacing in the time domain as $\Delta t'$, and

$$h = \Delta t / \Delta t' \quad (102)$$

Similarly, we will define $f_r' = 1/\Delta t'$, so that

$$h = f_r' / f_r \quad (103)$$

Furthermore, we will define the number of samples in one repetition interval in the frequency domain as

$$N_r' = f_r' / \Delta f \quad (104)$$

and as a result

$$N_r = N_r' / h \quad (105)$$

Interpolation in the time domain causes spurious responses in the frequency domain [8]. If we wish to hold these spurious responses to 50 dB below the desired clutter power, we must choose $f_r' \geq 10 f_{3dB}$ if linear interpolation is used [8, Eq. 8.61]. Now we can use (103) and write $h \geq 10 f_{3dB} / f_r$; however, if this ratio is greater than unity we will not implement interpolation. Thus we can write an equality that will provide -50 dB spurious responses for linear interpolation as

$$h = \min(10 f_{3dB} / f_r, 1) \quad (106)$$

If we use $\Delta f \leq 0.63 f_{3dB}$ in (104) we can also write

$$N_r' \geq 1.60 f_r' / f_{3dB} \quad (107)$$

which is equivalent to (101) applied to the reduced sampling rate. If we combine (107) with (103) we obtain

$$N_r' \geq 1.60 h \cdot f_r / f_{3dB} \quad (108)$$

and with (106) we have

$$N_r' \geq 16, \quad \text{if } f_{3dB} / f_r \leq 0.1 \quad (109a)$$

$$\geq 1.6 f_r / f_{3dB}, \quad \text{otherwise} \quad (109b)$$

Note that (109a) provides the largest lower bound on N_r' , so that we can write for all cases $N_r' \geq 16$. This is a factor of 20 less than N_r in the example in Section 4.1.

Let us define $N = T/\Delta t$. Then from (93), (94), and (97) we can write

$$N_r \geq N + f_r / f_{3dB} \quad (110)$$

With (105) and $h = 10 f_{3dB} / f_r$, we obtain another lower limit on N_r' as

$$N_r' \geq (N + 10)h \quad (111)$$

usually for a ground based surveillance radar this limit will be lower than (109a).

In Table 1 we list the performance that can be obtained for various parameter choices. The use of the table will begin with a specification of the desired spurious spectral response level and accuracy. For efficient computation the lowest value of N_r' should be used, although some consideration should be given to the use of $N_r' = 16, 32, 64, \text{etc.}$, because these values are especially efficient FFT sizes. We will be given N , the number of time samples being utilized, and the ratio f_{3dB} / f_r . We will then compute

$$h = (f_r' / f_{3dB}) f_{3dB} / f_r \quad (112)$$

where (f_r' / f_{3dB}) is given in the table. If $h > 1$ we will not implement interpolation. Next we will check the constraint on N in (111), which we rewrite as

$$N \leq N_r' / h - (f_r' / f_{3dB}) \quad (113)$$

Table 1. Performance for Various Interpolation Parameters

Option	N_r	f_r'/f 3dB	$\Delta f/f$ 3dB	Spurs [*]	Accuracy ^{**}	Constraint
A	16	10	0.625	-50 dB	2.0%	$N \leq 16/h-10$
B	20	10	0.50	-50	0.1	$\leq 20/h-10$
C	20	12	0.60	-53	1.5	$\leq 20/h-12$
D	24	12	0.50	-53	0.1	$\leq 24/h-12$
E	32	16	0.50	-58	0.1	$\leq 32/h-16$
F	40	16	0.40	-58	--	$\leq 40/h-16$
G	48	16	0.333	-58	--	$\leq 48/h-16$
H	64	16	0.25	-58	--	$\leq 64/h-16$

* Spurious spectral responses for linear interpolation [8, Eq. 8.61]

** Sampling error [8, Figure 8.5 for $N = 7$ and $m = 2$]

If it is satisfied then we can use that particular option; otherwise, we will have to use a larger value of N_r' . Let us now consider the example $f_{3dB}/f_r = .01$ and $N = 32$, and assume that Option A would be suitable if it satisfies the constraint on N . We obtain $h = 0.10$ and the constraint $N \leq 150$, which works. On the other hand, suppose that $f_{3dB}/f_r = .05$ and the performance of Option E were the minimum acceptable. Now we obtain $h = 0.80$ and $N \leq 64$, which works (Options F and G would also work). We should comment that when h is close to unity, as in the last example, interpolation may not offer any computational advantage over the straight FFT approach. We could have chosen $\Delta f = f_r/64 = .3125 f_{3dB}$, and we could have utilized $N = 64 - 20 = 44$ time samples without interpolation, which would have resulted in a slightly faster computation. Interpolation pays off when $h < 0.5$, especially so for very small values of h . When $0.5 \leq h < 1.0$ the issue is not so decisive as many factors come into play.

We can compare the computation times for the brute-force and FFT (with interpolation) approaches. For the former the computation time is proportional to $N \cdot N_r'$, where N_r' is the number of samples defining the power spectral density. For the FFT the computation time is proportional to $N_r' \log_2 N_r'$. On one particular computer the brute-force approach will be faster when N satisfies the following:

N_r'	$N_t = 5$	$N_t = 7$
16	$N \leq 5$	$N \leq 4$
32	$N \leq 14$	$N \leq 10$

For $N_r' = 16$ the constraints of $N \leq 4$ or $N \leq 5$ will not usually be of practical interest in radar simulations; even if we were interested in simulating a few samples it would be still relatively efficient to utilize the FFT approach with interpolation. For $N_r' = 32$ there might be a few situations in which the brute-force approach would offer a computational advantage, but the advantage will never be great.

5. SIMULATION PROGRAM

A very convenient, flexible, and computationally efficient Fortran program has been developed to simulate detection performance for a ground-based radar against airborne targets. The quadrature components of the video signal are simulated in the neighborhood of the target, and the received signal is processed in the receiver in the same manner and sequence as it would be in an actual radar. The scattering environment is described in a statistical manner, as well as the fine-scale location of the target within the azimuth beam, range gate, and Doppler filter. By repeating the run with independent random inputs, the detection statistics are accumulated in the form of cumulative probability versus threshold setting. If a target is absent, then the output is probability of false alarm versus threshold setting. Importance sampling techniques can be used to increase the efficiency of this calculation.

Practically any situation of interest can be simulated with parameters specified on input. Listings of the program as given in Appendix A include a description of all input parameters, as well as the procedures, assumptions, and limitations within each simulation step. The main program (MAIN) acts as a driver, in which calls are made to subroutines that generate specific signals such as

- the target (TARGET)
- clutter (CLUTTR)
- receiver noise (NOISE)

in addition to a subroutine that processes the received signal (PROCES) and accumulates detection performance statistics (DSTINL, DSTPNT, DSTFUN). The simulation procedure can be easily modified. For example, as written, the main program permits two types of clutter to be generated, ground plus rain or ground plus chaff; all three types can be generated with three calls to CLUTTR (which also requires additional parameters to be defined on input).

5.1 INPUT PARAMETERS

All input parameters are capable of being defined in DATA statements in the main program or via NAMELIST /VALUES/ for each simulation run. We will discuss these parameters as they relate to each simulation function.

The Target

The nominal location of the target is defined in terms of its range (R) and altitude above a spherical earth (HT). Its velocity is incorporated in the

simulation by defining in which Doppler filter (KFILT) the target appears (KFILT = 1 for dc). The average radar cross section (RCS) of the target is also specified, as well as a fluctuation parameter (KSWER). The fluctuation is assumed to be slow (pulse-to-pulse correlated but scan-to-scan independent), and the following cases can be handled:

<u>KSWER</u>	<u>Case</u>
0	nonfluctuating
1	Swerling Case 1
2	Swerling Case 3
>0	chi-square with 2*KSWER degrees of freedom

The target position as described above is nominal; for each statistical replication the actual position is varied randomly throughout the azimuth beam, the range gate, and the Doppler filter. The beam scanning and straddling losses are thus incorporated in the simulation.

Ground Clutter

Spatially nonhomogeneous ground clutter can be simulated. The model used in this program is a block correlated one, in which a value of σ_0 (the backscatter coefficient) is generated for one square of size RCOR from a statistical distribution that can be log-normal (SVP>0.) or Weibull (SVP<0.). The σ_0 for adjacent squares in the format of Figure 12 is generated independently, but within any square the σ_0 is constant (homogeneous) with an assumed Rayleigh amplitude distribution applying to each elemental area within the square. The global average σ_0 is defined as SIC0 on input. For homogeneous ground clutter set SVP=0.

Two models for the clutter spectrum (not including the scan modulation) are included. They are a Gaussian shape (IT=0) and a general shape (IT>0) of

$$s(f) = \frac{1}{1 + (f/f_0)^n}$$

where $n = IT$ (the main program restricts the choice to $IT = 2$ or 3). The half-power spectral width relative to the PRF is defined as FWTR. The above description for the clutter spectrum is designated as the "ac component." There is also a "dc component," the power of which is DCAC times the power in the ac component.

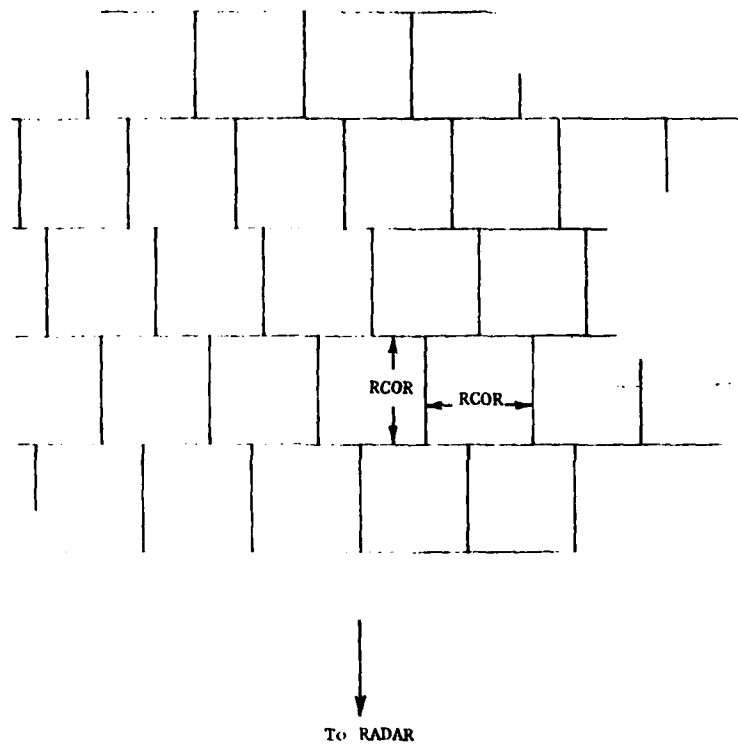


Figure 12. Format of Block Correlated Squares on the Ground
(there is random staggering from one row to the next)

The user must determine if there will be any ground clutter at all because of terrain masking, and whether there will be any enhancement of short range clutter in range-ambiguous situations ($R > C/(2 \cdot \text{PRF})$, where C is the propagation velocity).

Rain/Chaff Clutter

The volumetric clutter is assumed to be homogeneously distributed in a layer between two altitudes ($H1$ and $H2$, $H2 > H1$) relative to a spherical earth. The reflectivity, or backscatter cross section per unit volume, is defined as REF . For rain clutter set $H1=0$. A Gaussian shape is assumed for the fluctuation spectrum.

Wind

The spectral width of all types of clutter, and the mean Doppler of volumetric clutter, are functions of the wind speed (VW) and direction with respect to the antenna beam axis ($ANGL$).

The System

The radar system is defined by the following parameters:

- PT = peak transmit power
- $G\emptyset$ = peak antenna gain, one-way
- WL = wavelength
- BW = pulse (resolution) bandwidth
- PRF = pulse repetition frequency
- FL = product of noise figure and rf losses ($L > 0$)
- $AZ3DB$ = half-power width of azimuth beam, one way
- $ITYPE$ = type of scan (=1 for step, =2 for continuous)
- $AZDWEL$ = azimuth angle through which the antenna steps ($ITYPE=1$) or scans ($ITYPE=2$) from dwell to dwell*
- $PSAT$ = receiver saturation level, after pulse cancellation but before Doppler filtering
- $NCOH$ = number of pulses coherently integrated
- $NNCOH$ = number of pulse groups noncoherently integrated at same frequency
- $NCNCL$ = number of stages of pulse cancellation
- NR = number of range gates simulated (=1 for conventional threshold detection, >1 but odd for CFAR processing--the number of CFAR reference cells is $NR-1$)
- LAW = envelope detection law (=1 for linear, =2 for square law)

* a dwell consists of $NCOH \cdot NNCOH \cdot NCNCL$ pulses.

The processing stages in the receiver are assumed to be in the following order, beginning at the front end:

- range gating
- pulse cancellation
- saturation
- Doppler filtering
- envelope detection
- threshold detection or CFAR processing

The resultant clutter spectrum for the scanning antenna (ITYPE=2) is modeled on the basis of the Gaussian shape for the beam and input spectrum.

Simulation

The number of statistical replications under the same nominal conditions is specified as NREP. For false alarm analysis set RCS=0., and importance sampling can be invoked with KSW=1 (conventional sampling will be implemented with KSW=0). The distortion parameter for importance sampling is XM, which should be established by trial and error (too small of a value will result in infrequent false alarms, while too large of a value will cause underflow/overflow or other diagnostics; XM=2. is a reasonable first guess). As set by a DATA statement the number of azimuth samples simulated for clutter is NA=21, and the spacing of the samples is computed in MAIN as DA=.2*AZ3DB. The user can decrease the running time by choosing NA=11 and DA=.4*AZ3DB, or NA=9 and DA=.5*AZ3DB, to encompass ± 2 beam-widths of the azimuth mainbeam. Further out sidelobes can be accommodated by choosing NA and DA accordingly.

Debug printouts can be obtained by setting IDBG=0. In addition, it is possible to obtain a printout of the signal spectrum in the first range gate (or any range gate with suitable program modifications).

5.2 INTERPRETING THE RESULTS

For conventional threshold detection (non-CFAR) the output of the Doppler filter(s) under test is normalized by the input noise power divided by NCOH, which is the approximate (receiver) noise power in a Doppler filter on output (it would be the noise power if the filter weights were uniform). These detected outputs are then accumulated in a histogram, and the histogram is integrated to obtain the cumulative distribution function versus relative threshold setting. For CFAR processing the output of the Doppler filter under test is normalized by

the average of the NR-1 range gated outputs forming the CFAR reference (half of the gates are in front of the range gate under test, and half are behind it).

In order to determine what the threshold setting should be set $RCS=0$, to get the simulated false alarm probability versus threshold setting. Usually, the false alarm probability will be given. Then with $RCS>0$, we can read out the probability of detection that corresponds to the desired threshold setting.

5.3 EXAMPLES

If $NNCOH=1$, $KSWER=1$, $LAW=2$, and $NR=1$, all signals appearing in a non-clutter region of the Doppler spectrum will be Rayleigh amplitude and the probability of exceeding a threshold that is normalized by the average output noise power is given by

$$P(T) = e^{-T/(1 + SNR)} \quad (114)$$

where T is the normalized threshold and SNR is the output signal-to-noise ratio.

For the first example we will examine the false alarm performance with importance sampling. The simulated parameter values that deviate from the built-in data values are $RCS=0$, $SIG0=0$, $NCOH=1$, $KFILT=1$, $NREP=1000$, $KSW=1$, and $XM=13.8$. The simulation output is given in Table 2. Since $SNR=0$ in (114) we can easily compute what the theoretical results should be. For example, $P(8.0) = .000335$, $P(10.0) = .0000454$, and $P(12.0) = .00000614$; all values fall within 10% of the corresponding values in Table 2.

For the second example we will repeat the above conditions except we will utilize four Doppler filters ($NCOH=4$, $KFILT=3$). The simulation output is given in Table 3. Because of Hamming Weighting in the Doppler filtering process, the noise bandwidth of a Doppler filter is about 25% larger than the PRF divided by the number of filters ($NCOH$). If we thus multiply the values of T in Table 3 by 0.80 prior to the use of (114), we will get good statistical agreement.

For the third example, we add a target ($RCS=4$) so that the input signal-to noise ratio is a factor $NCOH$ higher ($4 \times 8.0 = 32.0$), reduced by the following effects

Antenna scan	.69	(-1.6 dB)
Doppler filter straddle	.90	(-0.5 dB)
Range gate straddle	.58	(-2.3 dB)
Total	.36	(-4.4 dB)

Table 2

THRESHOLD CROSSING STATISTICS

THRESHOLD T	FROM THAT DATA SET		
0.00	1.01380057	10.20	.00003933
0.00	.04770106	10.40	.00003326
0.00	.71950792	10.60	.00002719
0.00	.59084531	10.80	.00002040
0.00	.45420103	11.00	.00001635
1.00	.35374026	11.20	.00001351
1.00	.30390993	11.40	.00001198
1.00	.24210407	11.60	.00001074
1.00	.19800552	11.80	.00000886
1.00	.17010072	12.00	.00000652
2.00	.13804017	12.20	.00000539
2.00	.11601927	12.40	.00000464
2.00	.09399108	12.60	.00000413
2.00	.081002036	12.80	.00000339
2.00	.06050002	13.00	.00000273
3.00	.04640717	13.20	.00000221
3.00	.03630204	13.40	.00000185
3.00	.03100665	13.60	.00000120
3.00	.020100027	13.80	.00000086
3.00	.02130364	14.00	.00000073
4.00	.01740784	14.20	.00000064
4.00	.01350029	14.40	.00000057
4.00	.01120024	14.60	.00000047
4.00	.00911000	14.80	.00000039
4.00	.00774967	15.00	.00000031
5.00	.00630960	15.20	.00000025
5.00	.005004393	15.40	.00000022
5.00	.00390700	15.60	.00000017
5.00	.003100067	15.80	.00000014
5.00	.00250071	16.00	.00000012
6.00	.00200003	16.20	.00000010
6.00	.00170439	16.40	.00000008
6.00	.00122270	16.60	.00000007
6.00	.00142592	16.80	.00000006
6.00	.00117393	17.00	.00000005
7.00	.00091023	17.20	.00000003
7.00	.00070414	17.40	.00000002
7.00	.00062227	17.60	.00000002
7.00	.00046485	17.80	.00000001
7.00	.00042715	18.00	.00000001
8.00	.00031403	18.20	.00000001
8.00	.00020095	18.40	.00000001
8.00	.00020004	18.60	.00000001
8.00	.00015490	18.80	.00000001
8.00	.00013776	19.00	.00000000
9.00	.00011658	19.20	.00000000
9.00	.00009803	19.40	.00000000
9.00	.00007433	19.60	.00000000
9.00	.00005902	19.80	.00000000
9.00	.00004725	20.00	.00000000
10.00	.00003615		

Table 3

THRESHOLD CROSSING STATISTICS

THRESHOLD T	PROB THAT DATA.GE.T		
0.00	.18263636	10.20	.00039271
.20	.17811398	10.40	.00039266
.40	.17805438	10.60	.00021992
.60	.15218296	10.80	.00021918
.80	.15214634	11.00	.00021917
1.00	.15110649	11.20	.00021861
1.20	.15108545	11.40	.00006844
1.40	.15106213	11.60	.00006218
1.60	.14669242	11.80	.00005073
1.80	.14644394	12.00	.00005031
2.00	.14641651	12.20	.00005025
2.20	.13592534	12.40	.00002847
2.40	.13583332	12.60	.00002766
2.60	.12450066	12.80	.00002690
2.80	.09873398	13.00	.00002644
3.00	.09224525	13.20	.00002644
3.20	.06018307	13.40	.00002506
3.40	.06012936	13.60	.00002506
3.60	.06006562	13.80	.00002505
3.80	.05990731	14.00	.00002505
4.00	.05272144	14.20	.00002418
4.20	.04537115	14.40	.00002337
4.40	.03748572	14.60	.00002337
4.60	.01141535	14.80	.00002337
4.80	.01138476	15.00	.00002331
5.00	.01138207	15.20	.00000202
5.20	.01014165	15.40	.00000201
5.40	.01013885	15.60	.00000161
5.60	.01012378	15.80	.00000161
5.80	.00997911	16.00	.00000161
6.00	.00988891	16.20	.00000160
6.20	.00983875	16.40	.00000160
6.40	.00983671	16.60	.00000160
6.60	.00972320	16.80	.00000158
6.80	.00959998	17.00	.00000158
7.00	.00157811	17.20	.00000156
7.20	.00152927	17.40	.00000151
7.40	.00141015	17.60	.00000145
7.60	.00140948	17.80	.00000145
7.80	.00131221	18.00	.00000143
8.00	.00099402	18.20	.00000143
8.20	.00088224	18.40	.00000143
8.40	.00088083	18.60	.00000143
8.60	.00081441	18.80	.00000139
8.80	.00080644	19.00	.00000139
9.00	.00080610	19.20	.00000139
9.20	.00055735	19.40	.00000137
9.40	.00051021	19.60	.00000137
9.60	.00050608	19.80	.00000132
9.80	.00050546	20.00	.00000150
10.00	.00039346		

Thus the output signal-to-noise ratio is $SNR = 11.5$. In Table 4 we show the simulation results, which are in good statistical agreement with (114).

Table 4

THRESHOLD CROSSING STATISTICS

THRESHOLD T	PERCENT DATA, C.E.T		
0.20	1.00000000	10.20	.54000000
.20	.99000000	10.40	.54000000
.40	.97000000	10.60	.53000000
.60	.96000000	10.80	.52000000
.80	.95000000	11.00	.51000000
1.00	.94000000	11.20	.51000000
1.20	.93000000	11.40	.50000000
1.40	.92000000	11.60	.49000000
1.60	.91000000	11.80	.48000000
1.80	.90000000	12.00	.46000000
2.00	.89000000	12.20	.46000000
2.20	.88000000	12.40	.46000000
2.40	.87000000	12.60	.46000000
2.60	.86000000	12.80	.46000000
2.80	.85000000	13.00	.46000000
3.00	.84000000	13.20	.45000000
3.20	.83000000	13.40	.44000000
3.40	.82000000	13.60	.43000000
3.60	.81000000	13.80	.42000000
3.80	.80000000	14.00	.42000000
4.00	.79000000	14.20	.42000000
4.20	.78000000	14.40	.42000000
4.40	.77000000	14.60	.41000000
4.60	.76000000	14.80	.40000000
4.80	.75000000	15.00	.39000000
5.00	.74000000	15.20	.38000000
5.20	.73000000	15.40	.38000000
5.40	.72000000	15.60	.38000000
5.60	.71000000	15.80	.38000000
5.80	.70000000	16.00	.37000000
6.00	.69000000	16.20	.36000000
6.20	.68000000	16.40	.35000000
6.40	.67000000	16.60	.34000000
6.60	.66000000	16.80	.34000000
6.80	.65000000	17.00	.34000000
7.00	.64000000	17.20	.33000000
7.20	.63000000	17.40	.32000000
7.40	.62000000	17.60	.31000000
7.60	.61000000	17.80	.30000000
7.80	.60000000	18.00	.29000000
8.00	.59000000	18.20	.28000000
8.20	.58000000	18.40	.27000000
8.40	.57000000	18.60	.26000000
8.60	.56000000	18.80	.25000000
8.80	.55000000	19.00	.24000000
9.00	.54000000	19.20	.23000000
9.20	.53000000	19.40	.22000000
9.40	.52000000	19.60	.21000000
9.60	.51000000	19.80	.20000000
9.80	.50000000	20.00	.19000000
10.00	.49000000		

6. REFERENCES

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APPENDIX A
RADAR SIGNAL SIMULATION PROGRAM TO
DETERMINE DETECTION PERFORMANCE

FORTRAN LISTINGS

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*
*   RADAR SIGNAL SIMULATION PROGRAM
*   TO DETERMINE DETECTION PERFORMANCE
*
*   BY RL MITCHELL, MARK RESOURCES INC
*   PREPARED FOR WACC/HANSCOM AFB
*   SEPT 1980
*
*****

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C THE OBJECTIVE OF THIS SIMULATION PROGRAM IS TO DETERMINE DETECTION
C PERFORMANCE FOR A GROUND BASED RADAR AGAINST AN AIRBORNE TARGET. THE
C QUADRATURE VIDEO SIGNALS ARE GENERATED FOR THE TARGET, GROUND CLUTTER,
C RAINFALL CLUTTER, AND RECEIVER NOISE, AND THE COMBINED SIGNAL IS
C PROCESSED IN A SIMULATED RECEIVER THAT INCLUDES SOME OF ALL OF THE
C FOLLOWING OPERATIONS:..... PULSE CANCELLATION MTI, RECEIVER SATURATION,
C DOPPLER FILTERING VIA FFT, ENVELOPE DETECTION, NONCOHERENT INTEGRA-
C TION AND GILI-AVERAGING CLAR PROCESSING. IN THE ABSENCE OF A TARGET,
C IMPUTANCE SAMPLING CAN BE USED TO EFFICIENTLY AND RELIABLY ESTABLISH
C FALSE ALARM PERFORMANCE AT LOW FALSE ALARM RATES. THIS SIMULATION
C PROGRAM ALSO INCORPORATES MANY NOVEL FEATURES TO COVER A WIDE VARIETY
C OF ENVIRONMENTAL AND SYSTEM CONDITIONS.....

1. TABLE OF RECEIVING STATISTICS.

THE TANK IS ASSUMED TO BE SUBWAY FLUCTUATING COMPLETELY
CORRELATION FOR SUB-TANKS, BUT NON-TANK-TANK INDEPENDENT. THE
FLUCTUATING WAS SPECIFIED AS ONE OF THE CR-5000-1 MODELS THAT IN-
CLUDES FLUCTUATING AND SWERING CASES 1 AND 3. NO FREQUENCY
DEPENDENCY IS ASSUMED.

$$f_1 \in \mathcal{F}_1, f_2 \in \mathcal{F}_2, \dots, f_n \in \mathcal{F}_n, \text{ and } f_{n+1} \in \mathcal{F}_{n+1}.$$

THE JOINT COEFFICIENT IS ASSIGNED TO A BILCO CORRELATION. OTHER JOINT COEFFICIENTS IN THE GAUSSIAN COORDINATIONS MODE A RANDOM NUMBER IS GENERATED TO REPRESENT THE BACKSCATTER COEFFICIENT (SIGMA-ZERO) FOR THAT CELL. THESE CELLS WILL BE DESIGNATED AS CLUSTER CELLS. INDEPENDENT RANDOM NUMBERS ARE GENERATED FOR ALL OTHER CELLS. DIMENSIONS THEREIN CAN ALSO BE ASSIGNED.

U. N. GOVIL AND J. K. SHARMA: STATISTICS FOR CLOTHS CUTS.

THE RANGE NUMBERS ASSIGNED TO THE DUTY CELLS IN THE BACKGROUND PROCESS CAN BE GENERATED FROM A SINGLE BACKGROUND DISTRIBUTION FUNCTION. AN OVERALL AVERAGE BACKSCATTER COEFFICIENT IS SPECIFIED ON INPUT.

9. (b) $20000 - 10000 = 10000$ dollars.

THE SHAPE OF THE SECOND CLUSTER SPECTRUM CAN BE EITHER GAUSSIAN OR $1/(1+(f/f_0)^2)^{0.5}$. THE WIDTH OF THIS SPECTRUM IS PRE-

PORTIONAL TO THE WIND VELOCITY. THERE IS ALSO A CAPABILITY TO
ADD A DC COMPONENT THAT WOULD REPRESENT STATIONARY CLUTTER.

5. RAIN/CLUTTER CLUTTER.

HOMOGENEOUS VOLUMETRIC CLUTTER IS ASSUMED BETWEEN TWO ALTITUDES, THE LOWER ONE BEING AT GROUND LEVEL FOR RAIN. THE SPECTRUM FOR RAIN/CLUTTER CLUTTER IS GAUSSIAN SHAPED, WITH THE WIDTH BEING PROPORTIONAL TO THE WIND VELOCITY.....MANY FACTORS ENTER INTO THE CONSTANT OF PROPORTIONALITY AS DISCUSSED BY KATHANSEN (REF 2), AND ONLY A FIRST-CUT IS PROVIDED IN THIS PROGRAM.

6. ANTENNA SCANS.

TWO TYPES OF ANTENNA SCANS CAN BE SIMULATED. A CONTINUOUS SCAN OR A STEP SCAN. THE CONTINUOUS SCAN MODULATION FOR THE TARGET IS SIMULATED ON A PULSE-BY-PULSE BASIS, BUT FOR CLUTTER AN EQUIVALENT SPECTRAL BROADENING FACTOR IS COMPUTED ON THE BASIS OF AN ASSUMED GAUSSIAN-SHAPED AZIMUTH BEAM. THE ELEVATION BEAM IS ASSUMED TO BE A FAN BEAM, WITH A COSECANT-SQUARE SHAPE.

7. LOCATION OF TARGET.

THE TARGET IS ASSUMED TO BE RANDOMLY LOCATED WITHIN.....
A. THE AZIMUTH SCAN INCREMENT CORRESPONDING TO A DWELL
B. THE RANGE RESOLUTION CELL
C. A DOPPLER FILTER (IF DOPPLER FILTERING IS EMPLOYED)
OR THE PRT (IF NO DOPPLER FILTERING IS EMPLOYED).

8. DOPPLER SEARCH.

WHEN NO TARGET IS PRESENT (FALSE ALARM SIMULATION) A SEARCH IS PERFORMED OVER THOSE DOPPLER FILTERS THAT HAVE BEEN CALCULATED TO CONTAIN NO CLUTTER. FOR SOME LOWER-FREQUENCY RADARS, MAY BE, THE RAIN CLUTTER THAT APPEARS IN A DOPPLER FILTER MAY BE LESS THAN THERMAL NOISE SO THAT THE ALGORITHM USED IN THIS PROGRAM TO DETERMINE THE FIRST AND LAST FILTERS (K1 AND K2) WILL NOT BE CORRECT.....THE USER MUST ANALYZE THIS SITUATION, AS WELL AS THE SPECTRAL WIDTH OF RAIN CLUTTER.

THE FOLLOWING PARAMETERS ARE READ FROM NAMELIST /VALUES/.....

1. X = RANGE OF TARGET OR CELL UNDER TEST
2. Y = TARGET ALTITUDE
3. K1 = TARGET RCS (AVERAGE)
4. K2 = GROUND CLUTTER BACKSCATTER COEF (SIGMA-ZERO)
5. XPR = SPATIAL VARIABILITY PARAMETER
6. XPRVAR = CLUTTER CORRELATION DISTANCE
7. KPR = RATIO OF DC POWER TO AC POWER
8. KPRVOL = VOLUMETRIC CLUTTER (RAIN/CLUTTER) REFLECTION COEF (DBS PER UNIT VOLUME)

```

C      P1 = LOWER ALTITUDE OF RAIN/CHAFF CLUTTER (ZERO FOR RAIN)
C      P2 = UPPER ALTITUDE OF RAIN/CHAFF CLUTTER
C      WK = WIND SPEED
C      ANGL = ANGLE BETWEEN WIND AND ANTENNA BORESIGHT
C      PT = PEAK TRANSMIT POWER
C      CC = PEAK ANTENNA POWER GAIN (ONE-WAY)
C      WL = WAVELENGTH
C      BW = PULSE BANDWIDTH
C      PRF = PULSE REPETITION FREQUENCY
C      FL = PRODUCT OF NOISE FIGURE AND SYSTEM LOSSES
C      AZSEP = ONE-WAY HALF-POWER WIDTH (RAD)
C      AZSWEEP = AZ STEP BETWEEN DUELS (TYPE=1) OR
C              AZ SCAN DURING DUELS (TYPE=2)
C      PSAT = SATURATION POWER LEVEL
C      RM = DISTORTION OF SAMPLE MEAN FOR IMPORTANCE SAMP
C      NR = TOTAL NUMBER OF RANGE CELLS PROCESSED
C      NCCF = SIZE OF FFT IN DOPPLER PROCESSOR
C      NCCG = NUMBER OF PULSE GROUPS NONCOH INTEGRATED
C      NCCU = STAGES OF CANCELLATION
C      LAD = LAW OF FIRST DETECTOR
C      IT = PARAMETER SPECIFYING SPECTRAL SHAPE
C      KSWR = SWEEPING ONE-SQUARE PARAMETER (SEE REF.3)
C      NFILT = DOPPLER FILTER NUMBER IN WHICH TARGET APPEARS
C      NSRP = NUMBER OF REPETITIONS OF SIMULATION
C      RSW = SWITCH FOR IMPORTANCE SAMPLING (ON=1, OFF=0)
C      TYPE = TYPE OF SCAN (1 FOR STEP, 2 FOR CONTINUOUS)
C      DIM = DIMENSION OF X, Y, Z ARRAYS
C      TONG = DUELS FLAG (OFF DUELS=0, ON DUELS=GT.0)
C
C  THESE PARAMETERS ARE COMPUTED AND PRINTED IN NAMELIST /PARAM1/.....
C
C      WAMP = FIRST RANGE AMBIGUITY (=C/(2.*PRF))
C      LR = RANGE RESOLUTION CELL SIZE (=C/(2.*BW))
C      PSCL = SCALING FACTOR, CONVERTS RCS TO RECEIVED POWER
C      EL = TARGET ELEVATION ANGLE
C      EPWR = RECEIVED POWER FOR TARGET
C      CPWR = RECEIVED POWER FOR UNIT AREA OF GROUND CLUTTER
C      RPWR = RECEIVED POWER FOR UNIT AREA OF RAIN/CHAFF CLUTTER
C      SIGDEC = EQUIVALENT BACKSCATTER COEF FOR RAIN/CHAFF CLUTTER
C      PN = RECEIVER NOISE POWER
C      SRR = SINGLE-PULSE SIGNAL-TO-NOISE RATIO (RECEIVER NOISE ONLY)
C      SCR = SINGLE-PULSE SIGNAL-TO-CLUTTER RATIO (GROUND CLUTTER ONLY)
C      SVR = SINGLE-PULSE SIGNAL-TO-CLUTTER RATIO (VOLUME CLUTTER ONLY)
C      AREA = CLUTTER CELL AREA ON GROUND (= .707*R*DR*AZCH)
C      DAZ = AZIMUTH SCAN INCREMENT PER PULSE REPETITION PERIOD
C      AZRATE = ANTENNA SCAN RATE
C      UMAX = MAXIMUM DOPPLER OF WIND
C      UMAXR = UMAX/PRF
C      FDIR = FMAXR*COS(ANGLE)
C      FATHR = SPECTRAL WIDTH OF GROUND CLUTTER NORMALIZED TO PRF

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C   FWHMVL = SPECTRAL WIDTH OF RAIN/CLUTTER CLUTTER NORMALIZED TO PR
C   FWHMVK = SPECTRAL WIDENING OF CLUTTER DUE TO SCAN PCDULATION
C   INCLP = THRESHOLD NORMALIZATION (USED FOR CONVENTIONAL DETECTION)
C   NPULS = NUMBER OF PULSES PROCESSED FOR DWELL (=NCOH*ANCOH*NCNCL)
C   NTOT = NPULS*NS
C   K1 = FIRST FILTER FOR DOPPLER SEARCH                                PROCES
C   K2 = LAST FILTER FOR DOPPLER SEARCH                                PROCES
C
C   THESE PARAMETERS ARE SET IN DATA STATEMENTS.....
C
C   KRT = BOLTZMANN CONSTANT TIMES NOISE TEMPERATURE
C   C = PROPAGATION VELOCITY
C   RE = EFFECTIVE RADIUS OF EARTH (4/3 TIMES ACTUAL)
C   NA = NUMBER OF AZIMUTH SAMPLES COMPUTED FOR CLUTTER                CLUTTR
C   NT = NUMBER OF THRESHOLD SAMPLES
C   T1 = THRESHOLD VALUE FOR FIRST SAMPLE
C   DT = SPACING OF THRESHOLD SAMPLES
C   NS = NUMBER OF SPECTRAL SAMPLES IN SPECTRUM ANALYSIS
C
C   ALL UNITS MUST BE CONSISTENT THROUGHOUT.....
C
C       LENGTH = METERS
C       AREA = SQ. METERS
C       ANGLE = RADIANS
C       TIME = SEC
C       FREQ. = HZ
C       POWER = WATTS
C
C   IN THE FOLLOWING CHECK IS MADE FOR RANGE AMBIGUOUS OPERATION. IF R.GI.WAMP
C   THERE MAY BE ENHANCED CLUTTER DUE TO FOLDOVER OF SHORT-RANGE CLUTTER.
C   THIS ENHANCEMENT CAN BE ACCOMMODATED BY REDEFINING SIGC AND REF.
C
C   THERE IS NO PROVISION IN THIS PROGRAM FOR.....
C
C       1. PULSE COMPRESSION
C       2. FREQUENCY DIVERSITY
C       3. PULSE REPETITION INTERVAL STAGGERING
C
C   THE X(1),Y(1), AND Y1 ARRAYS MUST BE DIMENSIONED AS LARGE AS NTOT.
C   XARRAY Y1 MUST BE DIMENSIONED AS LARGE AS NT.
C   XARRAYS AND Y1 MUST BE DIMENSIONED AS LARGE AS NS.
C
C   THE FOLLOWING REFERENCES ARE NOTED.....
C
C       1. J. J. STEIN, "RADAR SIGNAL SIMULATION", ARTECH HOUSE, 1976
C       2. J. J. STEIN, "RADAR DESIGN PRINCIPLES", MC GRAW-HILL, 1968
C       3. J. J. STEIN, "RECURSIVE METHODS FOR .....", IEEE/TCS, JUL 71
C       4. J. J. STEIN, "STOCHASTIC HANDBOOK OF MATHEMATICAL FUNCTIONS
C
C   DIMENSION X(1250),Y(1250),Y1(250),Y1(250),S(128),PGT(128)

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DIMENSION T(1:101)
COMMON /ZICG/ 1000
COMMON /ISAP/ KSW,ISW,XY,NSUM,NSUM
COMMON /ZOM1/ NS,NPULS,NA,NCOM,F,DF,DA,AZ,3DH
COMMON /ZEPF/ KSWK(256)

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```

NAMELIST /VALUES/ R,H1,KCS,SIG0,SVP,RCUR,CCAC,KEF,H1,H2,VW,ANGL,
1 PT,GC,WL,FW,PRF,PL,AZ,3DB,AZDWEL,PSAT,XP,
2 NS,NCGH,NALGH,NCNCL,LAW,IT,KSWR,KFILT,NREP,KSA,
3 ITYPE,101M,100G
NAMELIST /PARAM1/ KAMP,DR,PSCL,EL,TPDW,CPW,RPCW,SIGGEO,PN,SNK,
1 SCH,SZZ,AREA,DAZ,AZRATE,FMAX,FMAXTR,FOTR,
2 FSTRGR,FSTRVL,FSTRSN,TAORM,NPULS,NTOT,K1,K2
NAMELIST /PARAM2/ DA,NA,KKT,C,PI,NT,TL,DT

```

```

DATA R/1.0/      *HT/1000.0/      *RCS/1.0/      *KSWR/1/
DATA PRF/1.0/    *PT/1.0/      *GC/4.0/      *WL/1/
DATA FW/1.0/     *PRF/1000.0/    *EL/10.0/     *PSAT/1.0/
DATA NA/1/       *NCOM/10/      *NCOM/1/      *NCNCL/0/
DATA LA/2.0/     *SIG0/1.0-4/    *SVP/0.0/     *RCUR/1.0/
DATA CCAC/0.0/   *IT/0/        *VW/10.0/     *ANGL/0.0/
DATA KSW/0.0/    *H1/0.0/      *H2/1000.0/   *NREP/1/
DATA KST/0.0-21/ *C/1.0/      *AZ 3DB/.05/   *AZDWEL/.05/
DATA ITYPE/1/    *101M/256/     *R/18.5/      *NA/21/
DATA NT/101/     *TL/0.0/      *100G/0/
DATA 100KPI/12.5/ *DT/.2/      *NS/128/

```

```

5 READ VALUES
IF(NREP.EQ.0) STOP
PRINT VALUES
NPULS=NCOM*NCOM*NCNCL
NT=NPULS*NS
KKT(1)=12.5
C IF(X(1),N0,125.0 STEP 30
IF(NCOM.EQ.0) STEP 31
IF(NTOT.GT.101M) STEP 32
IF(EL.GT.0.05,11.0,3) STEP 33
IF(RCUR.EQ.1.0,0.0,KFILT.GT.NCOM) STEP 34
IF(MOD(NS,2).EQ.0) STEP 35
IF(LAW.EQ.1.0,0.0,LAW.GT.2) STEP 36
IF(KSW.EQ.0.0,0.0,KSW.GT.1) STEP 37
C COMPUTE AZIMUTH SAMPLE SPACING
DA=.25A/1000
C COMPUTE PEAR SCALING IN RADAR RANGE EQUATION
PSCL=PI*CGO*EL**5.7741/0.0001/2000
C COMPUTE ELEVATION ANGLE OF TARGET
EL=PI/2-S/42.5-1)
C COMPUTE RECEIVED TARGET POWER
100*SG,

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```

      IF (EL.GE.0.) THEN PSC=PSCL*RCOS*ELGAIN(EL)
C  COMPUTE RECEIVED CLUTTER POWER FROM UNIT AREA ON GROUND (ZERO ELEVATION
C  TILT ANGLE IS ASSUMED)
      CPCL=PSC*SIGO*ELGAIN(0.)
C  COMPUTE EQUIVALENT SIGMA-ZERO OF RAIN/CHAFF
      DM=C1*(H2-H1)
      H=H1
      SUM=0.
      IF (EL.LE.0.) GO TO 22
      DO 20 J=1,100
        S=F/K-R/(2.*RI)
        IF (EL.GT.0.) SUM=SUM+ELGAIN(EL)
      20 CONTINUE
      22 SIGOCL=REF*SUM*DM
C  COMPUTE RECEIVED RAIN/CHAFF CLUTTER POWER PER UNIT CROSS-SECTION AREA
      RPLP=PSC*SIGOCL
C  COMPUTE NOISE POWER
      PN=XN1*RW*RI
C  COMPUTE SIGNAL-TO-NOISE RATIO (RECEIVER NOISE ONLY)
      SNR=1./PN
      IF (PK.GT.0.) SNR=TPCW/PN
C  COMPUTE RANGE CELL WIDTH
      RS=(Z/2.*RW)
C  COMPUTE FIRST RANGE AMBIGUITY
      SAMP=(Z/2.*RW)
C  COMPUTE CLUTTER CELL AREA ON GROUND
      A=RS*Z*RW*RW*RW*RW
C  COMPUTE SIGNAL-TO-CLUTTER RATIO (GROUND CLUTTER)
      SCW=1./RW
      IF (PK.GT.0.) SCW=TPCW/(RPLP*A)
C  COMPUTE SIGNAL-TO-CLUTTER RATIO (VOLUME CLUTTER)
      SVW=1./RW
      IF (PK.GT.0.) SVW=TPCW/(RPLP*A*RI)
C  COMPUTE REMAINING SCAN PARAMETERS
      DAZ=0.
      IF (TPCW.GT.1.) DAZ=AZRW/ELGAIN(PULS)
      AZRAT=DAZ*RW
C  COMPUTE WIND PARAMETERS
      FMAX=2.*RW*RI
      FMAX1=FMAX/200
      FWH=1./FMAX*0.1*FMAX
C  COMPUTE GROUND AND VOLUME CLUTTER SPECTRAL WIDTHS (SEE NATHANSONS BOOK
C  ON SECTION 5.2.9.1 AND 7.2.12 FOR MORE INFORMATION)
      FWHG=1000*FMAX1
      FWHV=1200*FMAX1
C  COMPUTE CLUTTER SPECTRAL BROADENING DUE TO SCAN MODULATION
      FWHM=1200*RW*RI*DAZ*RW
      FWHM=FWHM/200
C  COMPUTE SPECTRAL WIDTHS

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```

CALL ASUMINTOT,XR,YR,XR)
CALL ASUMINTOT,XI,YI,XI)
IF (IRG,LE,CI) GO TO 39
CALL PNTIYI,NTOT,-4.0H    YR,ZOHOUTPUT OF CLUTTER-RK )
CALL PNTIYI,NTOT,-4.0H    YI,ZOHOUTPUT OF CLUTTER-RK )
IF (S,LE,C) GO TO 38
N1=ASAZDILL*IRAN(0.1-1.1)
CALL TASCET(YR,YI,TCD,KSWR,KELT,AZL,DAZI)
CALL ASUMINTOT,XR,YR,XR)
CALL ASUMINTOT,XI,YI,XI)
IF (IRG,LE,CI) GO TO 38
CALL PNTIYI,NTOT,-4.0H    YR,ZOHOUTPUT OF TARGET )
CALL PNTIYI,NTOT,-4.0H    YI,ZOHOUTPUT OF TARGET )
C HERE WE PERFORM A SPECTRAL ANALYSIS ON THE FIRST RANGE RING
OR CALL XMITI-NS,0,YR)
CALL XMITI-NS,0,YI)
CALL XMITIIMPULS,XR,YR)
CALL XMITIIMPULS,XI,YI)
CALL SPCTI(YR,YI,IMPULS,NS,.08,WGT)
CALL ASUMNS,YR,S,S)
C HERE WE PROCESS THE SIGNAL
CALL PROCES(XR,XI,K1,K2,LAW,ANCDH,KCNCL,PSAT)
NK=K2-K1+1
ATL=
IF (NS,EC,1) WCONSUM=K1*XF(-CI,-1,ZM)*WSUM)
IF (ATL,0.1) GO TO 42
DO 40 K=1,NK
CALL PSTI(XR(XI-ZINORM,M))
40 CONTINUE
42 CONTINUE
C PRINT OUT SPECTRUM OF FIRST RANGE RING
SPCTI=1,ZSUMNS,S)
CALL PSN(-NS,SNRM,S,S)
CALL PNTIYI,NS,0.0H    S,ZOH SPECTRUM OF 1ST GATE)
C PRINT OUT THRESHOLD STATISTICS
CALL PSTI(GNCTI)
PRINT 110
DO 60 J=1,NT
T=1+J-10951
PSOP=1.-TTCJ)
PRINT 111,T,PSOP
60 CONTINUE
GO TO 5
65 PRINT 112,NSUM,WSUM,AL
STOP

110 C *****THRESHOLD CROSSING STATISTICS**
111 C *****THRESHOLD CROSSING DATA*****
112 C *****IMPORTANCE SAMPLING WEIGHT IS NEGATIVE**
113 C *****
114 C *****

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```

C IN THIS SUBROUTINE THE COMPLEX VIDEO SIGNAL FOR THE TARGET IS CREATED
C IN THE ARRAY=PAIR (XR,XI). THE TARGET SUPPLIER IS RANDOM WITHIN THE
C FILTER NUMBER KEIL (IF NCOM.GT.1) OR WITHIN THE PRF (IF NCOM=1).
C THE AZIMUTH BEAM SWEEPS PAST THE TARGET DURING THE DWELL OF NPULS
C PULSES BEGINNING AT AZ1 WITH AN AZIMUTH INCREMENT OF DAZ FOR EACH
C PULSE. THE RCS PARAMETERS ARE.....
C
C      TPBW = TARGET AVERAGE RECEIVED POWER (=PSCL*RCSC*ELGAIN**2)
C      KSWER = SWERLING CHI-SQUARE PARAMETER (SEE REF.3)
C
C THE FORMAT OF THE XR,XI-ARRAYS IS AS IF THEY WERE DIMENSIONED
C
C      XR(NPULS,NR)      XI(NPULS,NR)
C
C
C      DIMENSION XR(1),XI(1)
C      COMMON /COM1/ NR,NPULS,NA,NCOM
C      COMMON /PRG/ IPRG
C      DATA TWPB1/5.2531453/
C      PDW=TPBW
C      IF(KSWER.EQ.0) GO TO 22
C      PRGD=1.
C      DO 20 J=1,KSWER
C      PRGD=PRGD*PRAND(0.)
C20  CONTINUE
C      PRD=TPBW*(1-ALOG(PRGD))
C22  K=SECT(1)
C      R=PRAND(0.)
C      F=(KEIL*PRAND(0.)-1.5)/ELGAIN(NCOM)
C      L=(N-1)/2+NPULS*1
C      IF(TPBW.GT.0) PRINT 100,A,R,F
C      CALL XMITE(NR*NPULS,0.,XR)
C      CALL XMITE(NR*NPULS,0.,XI)
C      AZ=AZ1
C      DO 30 K=1,NPULS
C      ARG=TRDPI*PRF
C      SR=COS(ARG)
C      SI=SIN(ARG)
C      AMP=A*AZ*GAIN(AZ)
C      XR(1)=AMP*(1.-R)*SR
C      XI(1)=AMP*(1.-R)*SI
C      IF(NC.LE.1) GO TO 25
C      XR(1+NPULS)=AMP*R*SR
C      XI(1+NPULS)=AMP*R*SI
C25  L=L+1
C      AZ=AZ+DAZ
C30  CONTINUE
C      IF(I)G=1
C100  PRINT 11//Z10 TA=0.1      A,R,F.....F12.4,ZF12.4)
C      END

```

```

SUBROUTINE CLUTTR(XR,XI,CPCW,PCOR,SVF,IT,DCAC,FWTR,FQTR)
C
C IN THIS SUBROUTINE WE GENERATE THE COMPLEX VIDEO SIGNAL FOR CLUTTER
C WITH A STEEP SCANNING ANTENNA. THE CLUTTER SIGNAL IS OUTPUT TO THE
C ARRAY PAIR (XR,XI). THE PARAMETERS ARE.....
C
C   CPCW = CLUTTER RCS PER UNIT AREA ON THE GROUND SCALED TO RECEIVED
C           POWER (=SIGO*FSCL)
C   R = RANGE OF INTEREST
C   DR = RANGE RESOLUTION CELL SIZE
C   DA = AZIMUTH SAMPLE SPACING
C   NA = NUMBER OF AZIMUTH SAMPLES
C   RCCR = CLUTTER CORRELATION DISTANCE
C   SVP = SPATIAL VARIABILITY PARAMETER
C   FQTR = MEAN DOPPLER OF CLUTTER RELATIVE TO THE PRF
C   FWTR = DOPPLER WIDTH OF CLUTTER RELATIVE TO THE PRF
C   IT = PARAMETER SPECIFYING SPECTRAL SHAPE (SEE SUB. PARSEC)
C   DCAC = RATIO OF DC POWER TO AC POWER
C
C IN ADDITION N IS THE NUMBER OF SPECTRAL SAMPLES USED IN THE GENERATION
C OF THE CORRELATED RANDOM SEQUENCES. IT SHOULD BE AT LEAST AS LARGE
C AS NPULS+1./FWTR.
C
C THERE ARE TWO GENERAL CASES OF INTEREST.....
C
C           GROUND CLUTTER                RAIN/CHAFF CLUTTER
C           -----                -----
C   SVP=0. (NONHOMOGENEOUS)           SVP=0. (HOMOGENEOUS)
C   FQTR=0.                               IDTR=(2*VM/HL)*CCSI(ANGL)
C   DCAC=0.                               DC/AC=0.
C                                     IT=0
C
C IF THE ELEVATION BEAM DOES NOT HAVE PEAK GAIN IN THE GROUND AT THE
C RANGE OF INTEREST THEN CPCW MUST INCLUDE THE ELEVATION BEAM WEIGHTING.
C
C THE FORMAT OF THE XR,XI-ARRAYS IS AS IF THEY WERE DIMENSIONED
C
C           XR(NPULS,NR)      XI(NPULS,NR)
C
C   DIMENSION XR(1),XI(1)
C   COMMON /COM1/ NR,NPULS,NA,NCON,R,DR,DA
C   COMMON /ISAM/ KSW,ISW
C   COMMON /DRUG/ IDRG
C   DATA TWOPI/6.2831853/
C   NC=(NR+1)/2
C   R=CPCW*R*DR*DA
C   NR=NR+1
C   MR=-1
C   I=1
C   DO 40 I=1,NR

```

```

IF(MR.EQ.1) GO TO 25
SUM=0.
A7=-.5*(NA-1)*DA
AA=RANF(0.)
MA=0
W=SPTVAR(SVF)
DO 20 J=1,NA
IF(MA.EQ.1) GO TO 15
W=SPTVAR(SVP)
MA=AA
15 SUM=SUM+P*K*A7GAIN(A7)**2
A7=A7+CA
AA=AA+K*CA/RCOF
20 CONTINUE
MR=RF
25 PAV=SUM
IF(100*GT.0) PRINT 100,I,PAV
ISW=0
IF(KSW.EQ.1.AND.I.EQ.NC) ISW=1
CALL RANSEC(XR(L1),XI(L1),NPULS,FMT,IT,DCAC,PAV)
IF(F0TR.EQ.0.) GO TO 35
L=L1
DO 30 K=1,NPULS
ARG=TWCPI*F0TR*(K-1)
C=COS(ARG)
S=SIN(ARG)
YR=C*XF(L)-S*XT(L)
YI=C*XT(L)+S*XF(L)
XR(L)=YR
XI(L)=YI
L=L+1
30 CONTINUE
35 RF=RR+CA/RCOF
L1=L1+NPULS
40 CONTINUE
RETURN
100 FORMAT(///21H CLUTTR   I,PAV.....15,E12.4)
END

```

```

      SUBROUTINE NOISE(XR,XI,PN)
C
C   IN THIS SUBROUTINE THE THERMAL NOISE IS GENERATED IN THE ARRAY-PAIR
C   (XR,XI).
C
C       PN = AVERAGE NOISE POWER
C
C   THE FORMAT OF THE XR,XI-ARRAYS IS AS IF THEY WERE DIMENSIONED
C
C               XR(NPULS,NR)      XI(NPULS,NP)
C
      DIMENSION XR(1),XI(1)
      COMMON /ISAM/ KSW,ISW
      COMMON /COM1/ NR,NPULS
      NC=(NR+1)/2
      L=1
      DO 20 I=1,NR
        ISW=0
        IF(KSW.EQ.1.AND.I.EQ.NC) ISW=1
        DO 20 K=1,NPULS
          CALL GAUSS1(XR(L),XI(L),PN)
          L=L+1
20    CONTINUE
      RETURN
      END

```



```

SUBROUTINE PROCES(XF,XI,K1,K2,LAW,NCOCH,NCNCL,FSAT)
C
C IN THIS SUBROUTINE THE SIGNAL PROCESSING IN THE RECEIVER IS IMPLE-
C MENTED ON THE ARRAY-PAIR (XR,XI). THE ARGUMENTS ARE.....
C
C   K1 = FIRST FILTER FOR DOPPLER SEARCH
C   K2 = LAST FILTER FOR DOPPLER SEARCH
C   LAW = LAW OF FIRST DETECTOR
C   NCOCH = NUMBER OF PULSE GROUPS NONCOHERENTLY INTEGRATED
C   NCNCL = STAGES OF CANCELLATION
C   FSAT = SATURATION POWER LEVEL
C
C THE FOLLOWING CASES CAN BE HANDLED.....
C
C   PULSE CANCELLATION MTI WITH BINARY WEIGHTS
C
C   NCNCL = STAGES OF CANCELLATION (NCNCL+1 = NUMBER OF PULSES)
C   NCNCL = 0 CORRESPONDS TO NO PULSE CANCELLATION
C
C   DOPPLER FILTERING VIA FFT
C
C   NCOCH = SIZE OF FFT
C   NCOCH = 1 CORRESPONDS TO NO DOPPLER FILTERING
C
C   SATURATION BETWEEN PULSE CANCELLATION AND DOPPLER FILTERING
C
C   FSAT = SATURATION POWER LEVEL
C   FSAT = 1.E99 CORRESPONDS TO LINEAR PROCESSING AND SATURATION
C
C   LAW OF FIRST DETECTOR
C
C   LAW = 1 CORRESPONDS TO LINEAR DETECTOR
C   LAW = 2 CORRESPONDS TO SQUARE-LAW DETECTOR
C
C   NONCOHERENT INTEGRATION AT SAME CENTER FREQUENCY
C
C   NCOCH = NUMBER OF PULSE GROUPS NONCOHERENTLY INTEGRATED
C   NCOCH = 1 CORRESPONDS TO NO NONCOHERENT INTEGRATION
C
C   CFAR THRESHOLD REGULATION
C
C   NR = TOTAL NUMBER OF RANGE CELLS PROCESSED INCLUDING TARGET
C   NR-1 = NUMBER OF CFAR REFERENCE CELLS
C   NR = 1 CORRESPONDS TO CONVENTIONAL THRESHOLD PROCESSING
C
C THE INPUT COMPLEX SIGNAL IS IN THE ARRAY-PAIR (XF,XI), THE FORMAT OF
C WHICH IS AS IF THE DIMENSION WERE XR(INPULS,NR), XI(INPULS,NR).
C
C THE NONCOHERENTLY INTEGRATED OUTPUT OF FILTERS K1 THROUGH K2 APPEARS
C IN ARRAY XF IN SAMPLES 1 THROUGH K2-K1+1. IF NO DOPPLER FILTERING

```

C VIA FFT IS IMPLEMENTED THEN $NCCH=K1=K2=1$. USUALLY IF A TARGET IS
 C PRESENT THEN $K1=K2$ WILL CORRESPOND TO THE DOPPLER SAMPLE IN WHICH THE
 C TARGET WAS PLACED. IF NO TARGET IS PRESENT THEN THERE ARE SEVERAL
 C POSSIBILITIES.....

C 1. NO CLUTTER

C SET $K1=1$, $K2=NCCH$

C 2. GROUND CLUTTER ONLY WITH HANNING FILTER WEIGHTS

C SET $K1=3$, $K2=NCCH-1$

C 3. GROUND PLUS WEATHER CLUTTER

C SET $K1$ = FIRST DOPPLER SAMPLE NOT FLANKED
 C $K2$ = LAST DOPPLER SAMPLE NOT FLANKED

C FOR CFAR PROCESSING THE RANGE SAMPLE UNDER TEST IS ALWAYS THE CENTER
 C SAMPLE.

```

C
  DIMENSION XR(1),XI(1)
  DIMENSION Z(128)
  COMMON /COM1/ NR,NPULS,NA,NCCH
  COMMON /COM2/ IDRG
  IF (NPULS.NE.NCCH*NCCH+NCCL) STOP 1
  IF (MOD(NP,2).NE.1) STOP 2
  IF (NCCH.EQ.1.AND.K1.NE.1) STOP 3
  IF (NCCH.EQ.1.AND.K2.NE.1) STOP 4
  NREF=NP-1
  NP=NPULS
  DO 50 I=1,NR
    XI=(I-1)*NPULS+1
    IF (IDRG.LE.0.OR.I.GT.IDRG) GO TO 22
    CALL FNT(XR(1),NPULS,-4,6H XR,20HINPUT TO PROCES )
    CALL FNT(XI(1),NPULS,-4,6H XI,20HINPUT TO PROCES )
  22 IF (NCCL.EQ.0) GO TO 26
  C HERE WE IMPLEMENT PULSE CANCELLATION
  DO 25 M=1,NCCL
    L=1
    NP=NP-1
    DO 20 Y=1,NP
      XR(L)=XR(L)-XR(L+1)
      XI(L)=XI(L)-XI(L+1)
    L=L+1
  20 CONTINUE
  25 CONTINUE
  IF (IDRG.LE.0.OR.I.GT.IDRG) GO TO 26
  CALL FNT(XR(1),NP,-4,6H XR,20HAFTER PULSE CANCEL )
  CALL FNT(XI(1),NP,-4,6H XI,20HAFTER PULSE CANCEL )

```

```

26 IF (PSAT.GT.1.E90) GO TO 33
C HERE WE IMPLEMENT SATURATION
L=L1
DO 30 K=1,NP
P=XR(L)**2+XI(L)**2
IF (P.LE.PSAT) GO TO 30
A=SQRT(PSAT/P)
XR(L)=A*XR(L)
XI(L)=A*XI(L)
L=L+1
30 CONTINUE
IF (IDRG.LE.0.OR.I.GT.IDRG) GO TO 33
CALL FRNT(XR(L1),NF,-4,6H XR,20+AFTER SATURATION )
CALL FRNT(XI(L1),NF,-4,6H XI,20+AFTER SATURATION )
C HERE WE IMPLEMENT COHERENT FILTERING AND NONCOHERENT INTEGRATION
33 L=L1
CALL XFIT(-NCOH,0.,7)
DO 40 M=1,NCOH
CALL CCHFLT(XR(L),XI(L),LAH,NCNCL)
CALL ASUM(NCOH,XR(L),7,7)
L=L+NCOH
40 CONTINUE
IF (IDRG.LE.0.OR.I.GT.IDRG) GO TO 42
CALL FRNT(XR(L1),NF,-4,6H XR,20+AFTER CCH FILTER )
42 CALL XFIT(NCOH,7,XR(L1))
50 CONTINUE
KK=K2-K1+1
L1=M1
M1=M1+1
DO 55 I=1,NR
CALL XFIT(KK,XR(L1),XR(M1))
L1=L1+NPULS
M1=M1+KK
55 CONTINUE
IF (IDRG.LE.0) GO TO 57
CALL FRNT(XR,NF*KK,-4,6H XR,20+AFTER DETECTION )
57 IF (NR.EQ.1) GO TO 75
C HERE WE IMPLEMENT CFAR PROCESSING
NRC=(NF+1)/2
DO 70 K=1,KK
L=L1
SUM=0.
DO 60 I=1,NR
IF (I.NE.NRC) SUM=SUM+XR(L)
L=L+KK
60 CONTINUE
L=(NRC-1)*KK+K
XR(K)=NREF*XR(L)/SUM
70 CONTINUE
IF (IDRG.LE.0) GO TO 75
CALL FRNT(XR,KK,4,6H XR,20+AFTER CFAR )
75 RETURN
END

```

```

      SUBROUTINE CCHFLT(XR,XI,LAW,NCNCL)
C
C   IN THIS SUBROUTINE WE IMPLEMENT COHERENT FILTERING ON THE COMPLEX-VITEC
C   SIGNAL IN THE ARRAY-PAIR (XR,XI) OF LENGTH NCOH. THE DETECTED OUTPUT
C   APPEARS IN ARRAY XK OF LENGTH NCCH.
C
C   NCOH = SIZE OF FFT IN COFFLER FILTER BANK
C   NCNCL = NUMBER OF STAGES OF PULSE CANCELLATION
C   LAW = ENVELOPE DETECTION LAW (1=LINEAR, 2=SQUARE-LAW)
C
C   PULSE WEIGHTING IS APPLIED PRIOR TO COFFLER FILTERING. THE WEIGHTING
C   FUNCTION IS COSINE ON A RECTANGULAR, WHERE
C
C   ALPHA = 1. FOR UNIFORM WEIGHTING
C           = .08 FOR HANNING WEIGHTING
C
      DIMENSION XR(1),XI(1)
      DIMENSION C(128),W(128)
      COMMON /COM1/ NR,NPULS,NA,NCCH
      DATA ALPHA/.08/
      DATA C(1)/1./
      DATA N1,N2/0,0/,PI/3.14159265/
      IF(NCOH.EQ.1) GO TO 17
      IF(N1.EQ.NCOH.AND.N2.EQ.NCNCL) GO TO 15
      N1=NCOH
      N2=NCNCL
      CALL WEIGHT(K,NCOH,ALPHA)
      C(1)=1.
      GO TO 12 K=2,NCCH
      C(K)=SIN(PI*(K-1)/FLOAT(NCOH))**(2*NCNCL)
12 CONTINUE
15 CALL FFT1(NCOH,W,XR,XR)
   CALL PRGD(NCOH,W,XI,XI)
   CALL FFT2(XR,XI,NCCH,-1)
17 GO TO 25 K=1,NCCH
   XR(K)=XR(K)**2+XI(K)**2
   XRI(K)=XR(K)/C(K)
   IF(LAW.EQ.1) XR(K)=SQRT(XRI(K))
25 CONTINUE
   RETURN
   END

```

FUNCTION A7GAIN(A7)

C
C COMPUTES ONE-WAY POWER GAIN AT AZIMUTH ANGLE AZ (RAD). GAIN IS NORM-
C ALIZED TO A PEAK VALUE OF UNITY.

C
C IN THIS EXAMPLE A GAUSSIAN BEAM SHAPE IS ASSUMED WHERE

C
C A73DB = ONE-WAY HALF-POWER WIDTH (RAD)

C
C DIMENSION G(42)
C COMMON /COM1/ DUMMY(7),A73DB
C DATA NN/0/
C IF(NN.GT.0) GO TO 25
C NN=1
C GO 20 I=1,42
C G(I)=EXP(-(1.6651*(.05*(I-1))**2))
20 CONTINUE
25 A=20.*AES(A7)/A73DB
C IF(A.GT.40.) A=40.
C IA=A
C A=A-IA
C IA=IA+1
C A7GAIN=(1.-A)*G(IA)+A*G(IA+1)
C RETURN
C END

FUNCTION ELGAIN(EL)

C
C COMPUTES ONE-WAY POWER GAIN AT ELEVATION ANGLE EL (RAD). GAIN IS NORM-
C ALIZED TO A PEAK VALUE OF UNITY.

C
C IN THIS EXAMPLE A COSECANT-SQUARE BEAMSHAPE IS ASSUMED.
C

```

      DIMENSION G(42)
      DATA PI2/1.57079632/
      DATA EL0/.1/
      DATA NN/6/
      IF (NN.GT.0) GO TO 2
      NN=1
      DO 20 I=1,41
      F=PI2*(I-1)/40.
      G(I)=1.
      IF (E.GT.EL0) G(I)=(SIN(EL0)/SIN(E))**2
20  CONTINUE
      G(42)=G(41)
25  E=40.*EL/PI2
      IF (E.GT.40.) E=40.
      IF (E.LT.0.) E=0.
      IF =E
      E=E-IE
      IE=IE+1
      ELGAIN=(1.-E)*G(IE)+E*G(IE+1)
      RETURN
      END

```

```

      FUNCTION SPTVAR(SVP)
C
C THE RANDOM NUMBERS CORRESPONDING TO THE TERRAIN SPATIAL VARIABILITY
C ARE GENERATED IN THIS SUBROUTINE. THERE ARE THREE CASES.....
C
C      SVP.LE.0.      WEIBULL DISTRIBUTION, -SVP=WEIBULL PARAMETER
C      SVP.EQ.0.      HOMOGENEOUS TERRAIN
C      SVP.GT.0.      LOG-NORMAL DISTRIBUTION, SVP=STD DEV OF
C                      LOG-VARIATE (NEPERS)
C
C IN ALL CASES THE MEAN VALUE OF THE RANDOM NUMBER IS UNITY.
C
C SEE REF.1, EQS 9.10 AND 9.11
C
      DATA SS/0./
      SPTVAR=1.
      IF(SVP.EQ.0.) RETURN
      IF(SVP.GT.0.) GO TO 20
      IF(SS.EQ.SVP) GO TO 10
      SS=SVP
      A=-SVP
      GA=GAMMA(A+1.)
10  E=-ALOG(RANF(0.))
      SPTVAR=E**A/GA
      RETURN
20  G=GALSS(DUMMY)
      SPTVAR=EXP(SVP*(G-.5*SVP))
      RETURN
      END

```

SUBROUTINE RANSEQ(XR,XI,N,FWTR,IT,CCAC,PAV)

```

C
C IN THIS SUBROUTINE WE GENERATE A CORRELATED SEQUENCE OF GAUSSIAN
C RANDOM PHASORS IN THE ARRAY-PAIR (XF,XI) OF LENGTH N. THERE ARE TWO
C CASES FOR THE SPECTRAL SHAPE.....
C
C          IT=0          GAUSSIAN SPECTRAL SHAPE
C
C          IT.GT.0       SPECTRAL SHAPE IS 1/(1+(2F/FWTR)**IT)
C
C IN EITHER CASE FWTR IS THE RATIO OF THE 2-SIDED HALF-POWER SPECTRAL
C WIDTH AND THE PRF. THE ABOVE SPECTRUM IS DESIGNATED AS THE AC-COMPON-
C ENT. IT IS CENTERED AT DC. THERE IS ALSO A DC-COMPONENT ADDED TO THE
C ABOVE, WHERE CCAC IS THE RATIO OF DC TO AC POWER. THE AVERAGE POWER
C OF THE OUTPUT SAMPLES IS PAV.
C
C THE RANDOM SEQUENCE IS GENERATED BY THE FFT-METHOD WITH INTERPOLATION.
C THE PARAMETERS CHOSEN BELOW WILL PRODUCE A MAXIMUM OF -50DB SPURIOUS
C SPECTRAL RESPONSES WITH AN ACCURACY OF ABOUT 2-PERCENT.
C
  DIMENSION XR(1),XI(1),AR(257),AI(257),S(257)
  DATA FF/-1./,II/-1/,NN/-1/
  IF(FWTR.EQ.FF.AND.IT.EQ.II.AND.N.EC.NN) GO TO 35
  FF=FWTR
  II=IT
  NN=N
  IF(FWTR.LE.0.) GO TO 65
  H=AMIN1(10.*FWTR,1.)
  NF=16
10  NFH=NF/H-.5
  IF(NH.LE.NFH) GO TO 12
  NF=2*NF
  GO TO 10
12  IF(NF.GT.256) STOP 11
  NF2=NF/2+1
  LIM=3*NF*FWTR/H
  S(1)=1.
  DO 20 K=2,NF2
  S(K)=0.
  IF(K.GT.LIM) GO TO 15
  IF(IT.LE.0) S(K)=EXP(-(1.6651*(K-1)*H/(FWTR*NF))**.2)
  IF(IT.GT.0) S(K)=1./(1.+(2.*(K-1)*H/(FWTR*NF))**.IT)
15  S(NF+2-K)=S(K)
20  CONTINUE
  SUM=ZSUM(NF,S)
  S(1)=S(1)+CCAC*SUM
  SUM=SUM*(1.+CCAC)/PAV
  CALL FROD(-NF,1./SUM,S,S)
35  IF(FWTR.LE.0.) GO TO 65
  DO 40 K=1,NF

```



```

      CALL GAUSSI (AR(K),AI(K),S(K))
40  CONTINUE
      CALL FFT2 (AR,AI,NF,1)
      AR(NF+1)=AR(1)
      AI(NF+1)=AI(1)
      T=0.
      DO 50 K=1,N
        I=T
        TT=T-I
        I=I+1
        XR(K)=(1.-TT)*AR(I)+TT*AR(I+1)
        XI(K)=(1.-TT)*AI(I)+TT*AI(I+1)
        T=T+N
50  CONTINUE
      GO TO 70
65  CALL GAUSSI (PR,BI,FAV)
      CALL XMIT (-N,BR,XR)
      CALL XMIT (-N,BI,XI)
70  RETURN
      FNC

```

```

      SUBROUTINE GAUSSI(X,Y,P)
C
C  GENERATES RANDOM NUMBERS FOR THE FOLLOWING TWO DISTRIBUTION FUNCTIONS
C
C      CALL GAUSSI(X,Y,P)    GENERATES PAIR OF GAUSSIAN RANDOM
C                             PHASOR COMPONENTS OF AVERAGE POWER P
C
C      CALL EXPI(X)          GENERATES EXPONENTIAL RANDOM VARIABLE
C                             WITH UNIT AVERAGE POWER
C
C  THIS SUBROUTINE SET CAN ALSO BE USED WITH IMPORTANCE SAMPLING IF WE
C  SET.....
C
C      ISW = 1 TO ACTIVATE IMPORTANCE SAMPLING
C      XM = DISTORTION OF MEAN POWER
C
C  FOR THE FIRST CALL WHEN IMPORTANCE SAMPLING IS IN EFFECT WE MUST SET
C
C      NSUM = 0.
C      WSUM = 0
C
C  AFTER THE LAST CALL THE IMPORTANCE SAMPLING WEIGHT IS GIVEN BY
C
C      W = (XM**NSUM)*EXP(-(1.-1./XM)**NSUM)
C
C      COMMON /ISAM/ KSW,ISW,XM,NSUM,WSUM
C      DATA 770./
C      ITYPE=1
C      GO TO 10
C      ENTRY EXPI
C      ITYPE=2
10  XMM=1.
C      IF (ISW.EQ.1) XMM=XM
C      E=YMM*(-ALOG(RANF(7)))
C      X=E
C      IF (ISW.NE.1) GO TO 12
C      NSUM=NSUM+E
C      WSUM=WSUM+E
C      NSUM=NSUM+1
12  IF (ITYPE.GE.2) RETURN
C      E=SQRT(E*E)
15  A=RANF(7)
C      A=A+A-1.
C      B=RANF(7)
C      B=B+B-1.
C      A2=A*A
C      B2=B*B
C      C=A2+B2
C      IF (C.GT.1.) GO TO 15
C      X=E*(A2-B2)/C
C      Y=.5*E*A*B/C
C      RETURN
C      END

```

FUNCTION GAUSS(DUPHY)

C
C GENERATES GAUSSIAN RANDOM NUMBER OF ZERO MEAN, UNIT VARIANCE.
C

DATA I/O/
IF(I) 1,1,2
1 CALL GAUSS(A,E,2.)
GAUSS=A
I=1
GO TO 3
2 GAUSS=E
I=0
3 RETURN
END

```

      FUNCTION APROC(N,A,B,C)
C
C THIS SUBROUTINE PACKAGE PROCESSES ARRAYS. IN EVERY CASE IABS(N) IS
C THE LENGTH OF ALL ARRAYS. THERE ARE SEVERAL ENTRIES.....
C
C      CALL XMIT(-N,A,B)      B(K)=A(1) FOR K=1,N
C
C      CALL XMIT(N,A,B)      B(K)=A(K) FOR K=1,N
C
C      ZS=ZSUM(N,A)          ZS=SUM(A(K)) FOR K=1,N
C
C      CALL ASUM(-N,A,B,C)    C(K)=A(1)+B(K) FOR K=1,N
C
C      CALL ASUM(N,A,B,C)     C(K)=A(K)+B(K) FOR K=1,N
C
C      DD=DOT(N,A,B)          DD=SUM(A(K)*B(K)) FOR K=1,N
C
C      CALL PROD(-N,A,B,C)     C(K)=A(1)*B(K) FOR K=1,N
C
C      CALL PROD(N,A,B,C)      C(K)=A(K)*B(K) FOR K=1,N
C
C      E=ENGY(N,AR,AI)         E=SUM(AR(K)**2+AI(K)**2) FOR K=1,N
C
C      CALL PCWR(N,AR,AI,P)    P(K)=AR(K)**2+AI(K)**2 FOR K=1,N
C
C      CALL SWAP(N,A,B)        A(K) AND B(K) ARE SWAPPED FOR K=1,N
C
C SOME EXAMPLES ARE.....
C
C      SUP=ZSUM(N,A)           NORMALIZE A-ARRAY BY SUP
C      CALL FROG(-N,1./SUM,A,A)
C
C      E=ENGY(N,AR,AI)         NORMALIZE (AR,AI)-ARRAYS
C      ANCRN=1./SQRT(E)        BY TOTAL ENERGY
C      CALL PROD(-N,ANCRN,AR,AR)
C      CALL PROD(-N,ANCRN,AI,AI)
C
C      SUMSQ=DOT(N,A,A)        SUM-SQUARE OF ELEMENTS IN A-ARRAY
C
C NOTE THAT.....
C
C      E=ENGY(N,AR,AI)=DOT(N,AR,AR)+DOT(N,AI,AI)
C
C      DIMENSION A(1),E(1),C(1)
C      ENTRY PROC
C      IF(N) 10,12,15
C10  NN=-N
C      AA=A(1)
C      LL=11 K=1,NN
C      C(K)=AA*B(K)

```

```

11 CONTINUE
12 RETURN
15 DO 18 K=1,N
   C(K)=A(K)*B(K)
18 CONTINUE
   RETURN
   ENTRY ENGY
   EE=0.
   DO 25 K=1,N
     EE=EE+A(K)**2+B(K)**2
25 CONTINUE
   APRCC=EE
   RETURN
   ENTRY POWR
   CC 30 K=1,N
   C(K)=A(K)**2+B(K)**2
30 CONTINUE
   RETURN
   ENTRY ZSUM
   SUM=0.
   DO 35 K=1,N
     SUM=SUM+A(K)
35 CONTINUE
   APRCC=SUM
   RETURN
   ENTRY ASUM
   IF(N) 40,42,45
40 NN=-N
   AA=A(1)
   DO 41 K=1,NN
     C(K)=AA+B(K)
41 CONTINUE
42 RETURN
45 DO 48 K=1,N
   C(K)=A(K)+B(K)
48 CONTINUE
   RETURN
   ENTRY SWAP
   DO 50 K=1,N
     AA=A(K)
     A(K)=B(K)
     B(K)=AA
50 CONTINUE
   RETURN
   ENTRY DOT
   DP=0.
   DO 60 K=1,N
     DP=DP+A(K)*B(K)
60 CONTINUE
   APRCC=DP

```

```
RETURN  
ENTRY XM11  
IF (N) 62,66,67  
62 AA=-N  
AA=A(1)  
DO 65 K=1,NN  
E(K)=AA  
65 CONTINUE  
66 RETURN  
67 DO 70 K=1,N  
E(K)=A(K)  
70 CONTINUE  
RETURN  
END
```

FUNCTION GAMMA(Z)

C

C COMPUTES GAMMA FUNCTION OF A REAL ARGUMENT

C SEE REF.4, ECS 6.1.15, 6.1.16, AND 6.1.35.

C

```

      DATA A1/-.5748646/,A2/.9512363/,A3/-.6958588/,A4/.4245549/,
1      A5/-.1010678/
      X=Z-1.
      IF(X.LT.0.) STOP
      GAMMA=1.
10  IF(X.LT.1.) GO TO 20
      GAMMA=GAMMA*X
      X=X-1.
      GO TO 10
20  GAMMA=GAMMA*(((A5*X+A4)*X+A3)*X+Z)*X+A1)*X+1.)
      RETURN
      END

```

```

      SUBROUTINE PRNT(X,N,NPL,A,WHERE)
C THIS SUBROUTINE PRINTS OUT ARRAY-X OF LENGTH N. THE FORMAT IS.....
C
C      NPL.GT.0      F-FORMAT, NPL DECIMAL PLACES
C      NPL.LT.0      E-FORMAT, -NPL DECIMAL PLACES
C
C IN ADDITION.....
C
C      A = HOLLERITH NAME OF ARRAY BEING PRINTED (6H, RIGHT JUSTIFIED)
C      WHERE = HOLLERITH LABEL (20H)
C
      DIMENSION X(1),WHERE(2)
      PRINT 100,A,N,WHERE,A,(A,K,K=1,9)
      I1=1
      I2=MIN0(N,10)
10  NP=IARS(NPL)
      IF(NPL.GT.0) PRINT 101,I1,NP,(X(I),I=I1,I2)
      IF(NPL.LT.0) PRINT 102,I1,NP,(X(I),I=I1,I2)
      I1=I1+10
      I2=MIN0(I1+9,N)
      IF(I1.LE.N) GO TO 10
      RETURN
100 FORMAT(////13H PRINTOUT OF ,A6,8H(K),K=1,I4,FY2A10//
1      5H      K,6XA6,3H(K),9(1XA6,3H(K+,I1,1H)1/1)
101 FORMAT(I5,3X10F12.=)
102 FORMAT(I5,3X10E12.=)
      END

```



```

SUBROUTINE SPCTRMTXR,XI,NIN,NOUT,ALPHA,W)
C
C IN THIS SUBROUTINE WE COMPUTE THE POWER SPECTRUM OF THE COMPLEX TIME
C SEQUENCE IN THE ARRAY-PAIR (XR,XI) OF LENGTH NIN. THE POWER SPECTRUM
C IS RETURNED IN ARRAY XR, AND IT IS NOW OF LENGTH NOUT.
C
C THE SAMPLE SPACING OF THE POWER SPECTRUM IS 1/NOUT OF THE REPETITION
C FREQUENCY.
C
C A COSINE-ON-A-FEESTAL WEIGHTING IS APPLIED TO THE INPUT SAMPLES.
C ALPHA IS THE RATIO OF THE WEIGHTING FUNCTION AT THE EDGE TO THE
C CENTER. ALPHA=.08 FOR HAMMING AND ALPHA=1.0 FOR UNIFORM WEIGHTING.
C THE USER CAN SUPPLY HIS OWN WEIGHTING FUNCTION IN ARRAY W OF LENGTH
C NIN, BUT HE MUST SET ALPHA TO A NEGATIVE VALUE.
C
C ARRAY W IS A WORKING ARRAY AND IT MUST BE DIMENSIONED AS LARGE AS.....
C
C          NIN          IF NOUT.EQ.2**INTEGER
C          NIN+2*NOUT    IF NOUT.NE.2**INTEGER
C
C THE WEIGHTS ARE NORMALIZED SO THAT THE SUM IS UNITY.
C
      DIMENSION XR(1),XI(1),W(1)
      DATA NIN0/0/,NOUT0/0/
      IF(NIN.EQ.NIN0.AND.NOUT.EQ.NOUT0) GO TO 10
      IF(ALPHA.GE.0.) CALL WEIGHT(W,NIN,ALPHA)
      NIN0=NIN
      NOUT0=NOUT
10  N=NOUT-NIN
      IF(N) 40,15,12
12  CALL XMIT(-N,0.,XR(NIN+1))
      CALL XMIT(-N,0.,XI(NIN+1))
15  CALL PROD(NIN,W,XR,XR)
      CALL PROD(NIN,W,XI,XI)
      CALL FFT2(XR,XI,NOUT,-1)
      CALL PCWR(NOUT,XR,XI,XR)
      RETURN
40  PRINT 100,NIN,NOUT
      STOP
100 FORMAT(1X49H***ERROR IN SPECTRM, NOUT IS SMALLER THAN NIN ***/
1      /10X4HNIN=16,10X5HNOUT=16)
      END

```

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SIMULATING RADAR SIGNALS FOR DETECTION PERFORMANCE EVALUATION.(U)

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SUBROUTINE WEIGHT(W,N,ALPHA)

```

C
C THIS SUBROUTINE COMPUTES COSINE-ON-A-PEDESTAL WEIGHTS IN ARRAY W OF
C LENGTH N (N IS ALSO THE NUMBER OF PULSES). ALPHA IS THE RATIO OF THE
C WEIGHTING FUNCTION AT THE EDGE TO THE CENTER. ALPHA=.00 FOR PAMMING
C AND 1.0 FOR UNIFORM WEIGHTING. THE WEIGHTS ARE NORMALIZED SO THE SUM
C IS UNITY.
C
C IF THE WEIGHTS ARE TO BE USED AS PART OF A LONGER ARRAY OF LENGTH NR
C (WITH ZERO FILL), THE CALLING SEQUENCE IS.....
C
C      CALL XMIT(-NR,0.,W)
C      CALL WEIGHTS(W,N,ALPHA)
C
C TO CENTER THE WEIGHTS AT THE FIRST SAMPLE OF THE W-ARRAY (BEST DONE
C WHEN N IS ODD), FOLLOW THE ABOVE SEQUENCE WITH.....
C
C      CALL SHAFT(W,NP,-N/2)
C
C
C      DIMENSION W(1)
C      DATA TWOPI/6.2831853/
C      A=(1.+ALPHA)/2.
C      B=(1.-ALPHA)/2.
C      CN=(N+1)/2.
C      XN=N
C      DO 20 K=1,N
C      W(K)=A+B*COS(TWOPI*(K-CN)/XN)
20 CONTINUE
C      WNORM=1./7SUM(W)
C      CALL FFCO(-N,WNORM,W,W)
C      RETURN
C      END

```

```

      FUNCTION DEX(A,N)
C  COMPUTES DECIMAL VALUES.  THERE ARE TWO ENTRIES.....
C
C      D=DF(A)          D=10.*ALOG10(A)
C
C      CALL DRN(A,N)    A(K)=CB(A(K)), K=1,N
C
C  THE OUTPUT IS TRUNCATED TO THE INTERVAL -99, +99 OF.
C
      DIMENSION A(1)
      DE7(7)=10.*ALOG10 (AMAX1 (AMIN1(7.543E9,Z),1.2589E-10))
      ENTRY DR
      DRX=DR7(A(1))
      RETURN
      ENTRY DEX
      DO 20 I=1,N
      A(I)=DE7(A(I))
20  CONTINUE
      RETURN
      END

```

```

      SUBROUTINE DSTGEN(A,R,NN)
C     COMPUTES CUMULATIVE DISTRIBUTION FUNCTION OF SAMPLE DATA
C     TO INITIALIZE.....
C
C         CALL DSTINL(X1,XINC,N)           X1 = FIRST VALUE
C                                           XINC = VALUE INCREMENT
C                                           N  = NUMBER OF INCREMENTS
C
C     FOR EACH DATA POINT.....
C
C         CALL DSTPNT(X,W)                 X = DATA POINT VALUE
C                                           W = WEIGHT (USUALLY =1.)
C
C     TO COMPUTE THE DISTRIBUTION FUNCTION (AFTER LAST CALL TO DSTPNT).....
C
C         CALL DSTFUN(P)                   P = DISTRIBUTION FUNCTION
C
C     THE DISTRIBUTION FUNCTION IN ARRAY P IS.....
C
C         P(1) = FRCB(DATA,LT,X1)
C         P(2) = FRCB(DATA,LT,X1+XINC)
C         .
C         .
C         .
C         P(N) = FRCB(DATA,LT,X1+(N-1)*XINC)
C
C     ARRAY P IS DIMENSIONED FOR A MAX VALUE OF N=201
C
      DIMENSION A(1),P(202)
      ENTRY DSTINL
      N=NN
      IF(N.GT.201) STOP 22
      NS=0
      NP1=N+1
      X1=A(1)
      XINC=E
      CALL XPIT(-NP1,0.,P)
      RETURN
      ENTRY DSTFUN
      DO 10 L=2,NP1
      P(L)=P(L-1)+P(L)
10  CONTINUE
      DO 20 L=1,N
      P(L)=P(NP1)-P(L)
      A(L)=1.-P(L)/FLOAT(NS)
20  CONTINUE
C     PRINT 100,P(NP1)
100  FORMAT(16.6)

```

```
RETURN  
ENTRY DSTPNT  
NS=NS+1  
L=(A(1)-X1)/XINC+2.  
L=MAX0(L,1)  
L=MIN0(L,NF1)  
P(L)=P(L)+F  
RETURN  
END
```

```

SUBROUTINE FFT2YAR, AI, N, ISGN)
C
C THIS SUBROUTINE COMPUTES THE FFT IN THE ARRAY-PAIR (AR, AI) OF LENGTH N
C AND RETURNS THE RESULT IN THE SAME ARRAY PAIR.
C
C      ISGN = SIGN OF PHASE ARGUMENT IN FFT
C
C ARRAY W IN COMMON MUST BE DIMENSIONED AT LEAST AS LARGE AS 2*N IF
C N.NE.2**INTEGER.
C
  DIMENSION AR(1), AI(1)
  COMMON /TEMP/ W(1)
  I=1
  IF (ISGN.GT.0) I=0
  CALL FCURT(AR, AI, N, 1, I, 1, W(1), W(N+1))
  RETURN
  END

```

```

SUBROUTINE FOURT(DATAR,DATAI,NN,NDIM,IFRWD,ICPLX,WCRKE,WCRKI)
DIMENSION DATAR(1),DATAI(1),NN(1),WCRKE(1),WCRKI(1),IFACT(20)
C
C THE COOLEY-TUKEY FAST FOURIER TRANSFORM IN USASI BASIC FORTRAN
C
C EVALUATES COMPLEX FOURIER SERIES FOR COMPLEX OR REAL FUNCTIONS.
C THAT IS, IT COMPUTES
C  $FTFRAN(J1,J2,...)=SUM(DATA(I1,I2,...)*W1^{*(I1-1)*J1-1}$ 
C  $*W2^{*(I2-1)*J2-1}*...)$ 
C WHERE  $W1=EXP(-2*PI*SQRT(1-1)/NN(1))$ ,  $W2=EXP(-2*PI*SQRT(1-1)/NN(2))$ ,
C ETC. AND I1 AND J1 RUN FROM 1 TO NN(1), I2 AND J2 RUN FROM 1 TO
C NN(2), ETC. THERE IS NO LIMIT ON THE DIMENSIONALITY (NUMBER OF
C SUBSCRIPTS) OF THE ARRAY OF DATA. THE PROGRAM WILL PERFORM
C A THREE-DIMENSIONAL FOURIER TRANSFORM AS EASILY AS A ONE-DIMEN-
C SIONAL ONE, TWO IN A PROPORTIONATELY GREATER TIME. AN INVERSE
C TRANSFORM CAN BE PERFORMED, IN WHICH THE SIGN IN THE EXPONENTIALS
C IS +, INSTEAD OF -. IF AN INVERSE TRANSFORM IS PERFORMED UPON
C AN ARRAY OF TRANSFORMED DATA, THE ORIGINAL DATA WILL REAPPEAR,
C MULTIPLIED BY  $NN(1)*NN(2)*...$ . THE ARRAY OF INPUT DATA MAY BE
C REAL OR COMPLEX, AT THE PROGRAMMER'S OPTION, WITH A SAVING OF
C ABOUT THIRTY PER CENT IN RUNNING TIME FOR REAL OVER COMPLEX.
C (FOR FASTEST TRANSFORM OF REAL DATA, NN(1) SHOULD BE EVEN.)
C THE TRANSFORM VALUES ARE ALWAYS COMPLEX, AND ARE RETURNED IN THE
C ORIGINAL ARRAY OF DATA, REPLACING THE INPUT DATA. THE LENGTH
C OF EACH DIMENSION OF THE DATA ARRAY MAY BE ANY INTEGER. THE
C PROGRAM RUNS FASTER ON COMPOSITE INTEGERS THAN ON PRIMES, AND IS
C PARTICULARLY FAST ON NUMBERS RICH IN FACTORS OF TWO.
C
C TIMING IS IN FACT GIVEN BY THE FOLLOWING FORMULA. LET NTOT BE THE
C TOTAL NUMBER OF POINTS (REAL OR COMPLEX) IN THE DATA ARRAY, THAT
C IS,  $NTOT=NN(1)*NN(2)*...$ . DECOMPOSE NTOT INTO ITS PRIME FACTORS,
C SUCH AS  $2^{*K2} * 3^{*K3} * 5^{*K5} * ...$ . LET SUM2 BE THE SUM OF ALL
C THE FACTORS OF TWO IN NTOT, THAT IS,  $SUM2 = 2^{*K2}$ . LET SUMF BE
C THE SUM OF ALL OTHER FACTORS OF NTOT, THAT IS,  $SUMF = 3^{*K3} + 5^{*K5} + ...$ .
C THE TIME TAKEN BY A MULTIDIMENSIONAL TRANSFORM ON THESE NTOT DATA
C IS  $T = T0 + T1*NTOT + T2*NTOT*SUM2 + T3*NTOT*SUMF$ . FOR THE PRE-
C SENTImplementation FORTRAN 32 ON THE CDC 3300 (FLOATING POINT
C ADD TIME = SIX MICROSECONDS),
C  $T = 3000 + 600*NTOT + 50*NTOT*SUM2 + 175*NTOT*SUMF$  MICROSECONDS
C ON COMPLEX DATA.
C
C IMPLEMENTATION OF THE DEFINITION BY SUMMATION WILL RUN IN A TIME
C PROPORTIONAL TO  $NTOT^{*2}$ . FOR HIGHLY COMPOSITE NTOT, THE SAVINGS
C OFFERED BY COOLEY-TUKEY CAN BE DRAMATIC. A MATRIX 100 BY 100 WILL
C BE TRANSFORMED IN TIME PROPORTIONAL TO  $10000*(2+2+2+5+5+5+5) =$ 
C  $240,000$  (ASSUMING T2 AND T3 TO BE ROUGHLY COMPARABLE) VERSUS
C  $10000^{*2} = 1,000,000$  FOR THE STRAIGHTFORWARD TECHNIQUE.
C
C THE COOLEY-TUKEY ALGORITHM PLACES TWO RESTRICTIONS UPON THE
C NATURE OF THE DATA BEYOND THE USUAL RESTRICTION THAT

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0123
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0125
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```

C THE DATA FROM ONE CYCLE OF A PERIODIC FUNCTION. THEY ARE--
C 1. THE NUMBER OF INPUT DATA AND THE NUMBER OF TRANSFORM VALUES
C MUST BE THE SAME.
C 2. CONSIDERING THE DATA TO BE IN THE TIME DOMAIN,
C THEY MUST BE EQUI-SPACED AT INTERVALS OF DT. FURTHER, THE TRANS-
C FORM VALUES, CONSIDERED TO BE IN FREQUENCY SPACE, WILL BE EQUI-
C SPACED FROM 0 TO  $2\pi \cdot (NN(1)-1)/(NN(1) \cdot DT)$  AT INTERVALS OF
C  $2\pi/(NN(1) \cdot DT)$  FOR EACH DIMENSION OF LENGTH NN(I). OF COURSE,
C DT NEED NOT BE THE SAME FOR EVERY DIMENSION.
C
C THE CALLING SEQUENCE IS--
C CALL FCURT(DATAR,DATAI,NN,NDIM,IFRWD,ICPLX,WORKR,WORKI)
C
C DATAR AND DATAI ARE THE ARRAYS USED TO HOLD THE REAL AND IMAGINARY
C PARTS OF THE INPUT DATA ON INPUT AND THE TRANSFORM VALUES ON
C OUTPUT. THEY ARE FLOATING POINT ARRAYS, MULTIDIMENSIONAL WITH
C IDENTICAL DIMENSIONALITY AND EXTENT. THE EXTENT OF EACH DIMENSION
C IS GIVEN IN THE INTEGER ARRAY NN, OF LENGTH NDIM. THAT IS,
C NDIM IS THE DIMENSIONALITY OF THE ARRAYS DATAR AND DATAI.
C IFRWD IS AN INTEGER USED TO INDICATE THE DIRECTION OF THE FOURIER
C TRANSFORM. IT IS NON-ZERO TO INDICATE A FORWARD TRANSFORM
C (EXPONENTIAL SIGN IS -) AND ZERO TO INDICATE AN INVERSE TRANSFORM
C (SIGN IS +). ICPLX IS AN INTEGER TO INDICATE WHETHER THE DATA
C ARE REAL OR COMPLEX. IT IS NON-ZERO FOR COMPLEX, ZERO FOR REAL.
C IF IT IS ZERO (REAL) THE CONTENTS OF ARRAY DATAI WILL BE ASSUMED
C TO BE ZERO, AND NEED NOT BE EXPLICITLY SET TO ZERO. AS EXPLAINED
C ABOVE, THE TRANSFORM RESULTS ARE ALWAYS COMPLEX AND ARE STORED
C IN DATAR AND DATAI ON RETURN. WORKR AND WORKI ARE ARRAYS USED
C FOR WORKING STORAGE. THEY ARE NOT NECESSARY IF ALL THE DIMENSIONS
C OF THE DATA ARE POWERS OF TWO. IN THIS CASE, THE ARRAYS MAY BE
C REPLACED BY THE NUMBER 0 IN THE CALLING SEQUENCE. THUS, USE OF
C POWERS OF TWO CAN FREE A GOOD DEAL OF STORAGE. IF ANY DIMENSION
C IS NOT A POWER OF TWO, THESE ARRAYS MUST BE SUPPLIED. THEY ARE
C FLOATING POINT, ONE DIMENSIONAL OF LENGTH EQUAL TO THE LARGEST
C ARRAY DIMENSION, THAT IS, TO THE LARGEST VALUE OF NN(I).
C WORKR AND WORKI, IF SUPPLIED, MUST NOT BE THE SAME ARRAYS AS DATAR
C OR DATAI. ALL SUBSCRIPTS OF ALL ARRAYS BEGIN AT 1.
C
C EXAMPLE 1. THREE-DIMENSIONAL FORWARD FOURIER TRANSFORM OF A
C COMPLEX ARRAY DIMENSIONED 100 BY 16 BY 13.
C DIMENSION DATAR(100,16,13),DATAI(100,16,13),WORKR(100),WORKI(100)
C DIMENSION NN(3)
C NN(1)=100
C NN(2)=16
C NN(3)=13
C CALL FCURT(DATAR,DATAI,NN,3,1,1,WORKR,WORKI)
C
C EXAMPLE 2. ONE-DIMENSIONAL FORWARD TRANSFORM OF A REAL ARRAY OF
C LENGTH 64.
C DIMENSION DATAR(64),DATAI(64)

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```

C      CALL FOURTIGATAR,DATA1,64,1,1,0,0,0)
C
C      THERE ARE NO FROPP MESSAGES OR ERROR HALTS IN THIS PROGRAM. THE
C      PROGRAM RETURNS IMMEDIATELY IF NDT (OR ANY NN(I)) IS LESS THAN ONE.
C
C      THE SINE AND COSINE VALUES REQUIRED FOR THE TRANSFORM ARE
C      GENERATED RECURSIVELY. IF DOUBLE PRECISION IS AVAILABLE, IT IS
C      STRONGLY URGED THAT THE FOLLOWING VARIABLES BE SO DECLARED TO
C      REDUCE ACCUMULATION OF ROUND-OFF ERROR--
C      DOUBLE PRECISION TWOPI,THE1A,WSTPR,WSTPI,WPIPF,WHINI,WR,WI,WTEMP
C      * ,THEIM,WPIST,WPISTI,TWOWR,SR,SI,CLOSR,CLOSI,STMPF,STMPI
C      IN ADDITION, TWOPI SHOULD BE ASSIGNED A SUFFICIENTLY PRECISE
C      VALUE AND THE VARIOUS CALLS TO THE FUNCTIONS CCS AND SIN
C      SHOULD BE CHANGED TO DCOS AND DSIN.
C
C      PROGRAM BY NORMAN BRENNER FROM THE BASIC ALGORITHM BY CHARLES
C      RADER (BOTH OF MIT LINCOLN LABORATORY), MAY 1967. THE IDEA
C      FOR THE BIT REVERSAL WAS SUGGESTED BY RALPH ALTER (ALSC MIT LL).
C      ADAPTED FROM THE WORK OF JAMES W. COOLEY AND JOHN W. TUKEY,
C      AN ALGORITHM FOR THE MACHINE CALCULATION OF COMPLEX FOURIER
C      SERIES, MATH. COMPUT. 19, 98 (APRIL 1965), 247-301.
C
C      IF(NDT-1)GOTO,1,1
C      NTOT=1
C      DO 2 TDIM=1,NDIM
C      NTOT=NTOT*NN(IIDIM)
C      TWOPI=6.283185307
C
C      MAIN LOOP FOR EACH DIMENSION
C
C      NP1=1
C      DO 910 IIDIM=1,NDIM
C      N=NN(IIDIM)
C      NP2=NP1*N
C      IF(N-1)GOTO,900,5
C
C      IS N A POWER OF TWO AND IF NOT, WHAT ARE ITS FACTORS
C
C      M=N
C      NTWO=NP1
C      IF=1
C      IDIV=2
C      10  IQUOT=M/IDIV
C      IF(M-IDIV*IQUOT)50,11,11
C      IF(IIDIM)20,12,20
C      11  NTWO=NTWO*NTWO
C      IFACT(IF)=IDIV
C      IF=IF*1
C      M=IQUOT
C      12

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0201
0202
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	GO TO 10	0250
20	IDIV=3	0251
	INCN2=IF	0252
30	IQUOT=M/IDIV	0253
	IREM=M-IDIV*IQUOT	0254
	IF(IQUOT-IDIV)60,31,31	0255
31	IF(IREM)40,32,40	0256
32	IFACT(IIF)=IDIV	0257
	IF=IF+1	0258
	M=IQUOT	0259
	GO TO 30	0260
40	IDIV=IDIV+2	0261
	GO TO 30	0262
50	INCN2=IF	0263
	IF(IREM)60,51,60	0264
51	NTWO=NTWO+NTWO	0265
	GO TO 70	0266
60	IFACT(IIF)=M	0267
70	NONZF=NP2/NTWO	0268
C		0269
C	SEPARATE FOUR CASES--	0270
C	1. COMPLEX TRANSFORM	0271
C	2. REAL TRANSFORM FOR THE 2ND, 3RD, ETC. DIMENSION. METHOD--	0272
C	TRANSFORM HALF THE DATA, SUPPLYING THE OTHER HALF BY CON-	0273
C	JUGATE SYMMETRY.	0274
C	3. REAL TRANSFORM FOR THE 1ST DIMENSION, N ODD. METHOD--	0275
C	SET THE IMAGINARY PARTS TO ZERO.	0276
C	4. REAL TRANSFORM FOR THE 1ST DIMENSION, N EVEN. METHOD--	0277
C	TRANSFORM A COMPLEX ARRAY OF LENGTH N/2 WHOSE REAL PARTS	0278
C	ARE THE EVEN-NUMBERED REAL VALUES AND WHOSE IMAGINARY PARTS	0279
C	ARE THE ODD-NUMBERED REAL VALUES. UNSCRAMBLE AND SUPPLY	0280
C	THE SECOND HALF BY CONJUGATE SYMMETRY.	0281
C		0282
	ICASE=1	0283
	IFMTA=1	0284
	IF(100*FX)100,71,100	0285
71	ICASE=2	0286
	IF(100*FX)100,72,100	0287
72	ICASE=3	0288
	IF(100*FX)100,100,77	0289
73	ICASE=4	0290
	IFMTA=2	0291
	NTWO=NTWO/2	0292
	N=N/2	0293
	NP2=NP2/2	0294
	NTOT=NTOT/2	0295
	T=1	0296
	DO 80 J=1,NTOT	0297
	DATAF(J)=DATAF(I)	0298
	DATAF(J)=DATAF(I+1)	0299

80	I=I+2	0300
C		0301
C	SHUFFLE DATA BY BIT REVERSAL, SINCE $N=2^*K$. AS THE SHUFFLING	0302
C	CAN BE DONE BY SIMPLE INTERCHANGE, NO WORKING ARRAY IS NEEDED	0303
C		0304
100	IF (NONZF-1) 101,101,200	0305
101	NP2HF=NP2/2	0306
	J=1	0307
	DO 150 I2=1, NP2, NP1	0308
	IF (J-I2) 121, 130, 130	0309
121	I1MAX=I2+NP1-1	0310
	GO 125 I1=I2, I1MAX	0311
	DO 125 I3=I1, NTOT, NP2	0312
	J2=J+I3-I2	0313
	TEMPR=DATA(I3)	0314
	TEMPJ=DATA(J3)	0315
	DATA(I3)=DATA(J3)	0316
	DATA(J3)=DATA(I3)	0317
	DATA(I3)=TEMPR	0318
125	DATA(J3)=TEMPJ	0319
130	M=NP2HF	0320
140	IF (J-M) 150, 150, 141	0321
141	J=J-M	0322
	M=M/2	0323
	IF (M-NF1) 150, 140, 140	0324
150	J=J+M	0325
	GO TO 360	0326
C		0327
C	SHUFFLE DATA BY DIGIT REVERSAL FOR GENERAL N	0328
C		0329
200	DO 270 I1=1, NP1	0330
	DO 270 I3=I1, NTOT, NP2	0331
	J=I3	0332
	DO 260 I=1, N	0333
	IF (ICASE-3) 210, 220, 210	0334
210	WORKR(I)=DATA(IJ)	0335
	WORKI(I)=DATA(IJ)	0336
	GO TO 240	0337
220	WORKR(I)=DATA(IJ)	0338
	WORKI(I)=0.	0339
240	IFF2=NP2	0340
	IF=IFMTN	0341
250	IFF1=IFF2/IFACT(1F)	0342
	J=J+IFF1	0343
	IF (J-1)-IFF2) 260, 255, 255	0344
255	J=J-IFF2	0345
	IFF2=IFF1	0346
	IF=IF+1	0347
	IF (IFF2-NP1) 260, 260, 260	0348
260	CONTINUE	0349

	I2MAX=I3+NF2-NF1	0350
	I=1	0351
	DO 270 I2=I3,I2MAX,NP1	0352
	DATAI(I2)=WCRK(I1)	0353
	DATAI(I2)=WCRK(I1)	0354
270	I=I+1	0355
C		0356
C	SPECIAL CASE-- W=1	0357
C		0358
300	I1FNG=NF1	0359
	GO TO(302,301,302,302),ICASE	0360
301	I1FNG=NF0*(1+NFREV/2)	0361
302	IF(NTWC-NP1)600,600,303	0362
303	GO 430 I1=1,I1FNG	0363
	IMIN=NP1+I1	0364
	ISTEP=2*NP1	0365
	GO TO 330	0366
310	J=I1	0367
	DO 320 I=IMIN,NTOT,ISTEP	0368
	TEMPF=DATAI(I)	0369
	TEMPJ=DATAI(J)	0370
	DATAI(I)=DATAI(J)-TEMPF	0371
	DATAI(J)=DATAI(J)-TEMPF	0372
	DATAI(J)=DATAI(J)+TEMPF	0373
	DATAI(J)=DATAI(J)+TEMPF	0374
320	J=J+ISTEP	0375
	IMIN=IMIN+IMIN-I1	0376
	ISTEP=ISTEP+ISTEP	0377
330	IF(ISTEP-NTWC)310,310,331	0378
C		0379
C	SPECIAL CASE-- W=-SGRT(I-1)	0380
C		0381
331	IMIN=3*NP1+I1	0382
	ISTEP=4*NP1	0383
	GO TO 420	0384
400	J=IMIN-ISTEP/2	0385
	DO 410 I=IMIN,NTOT,ISTEP	0386
	IF(TEMPC)401,402,401	0387
401	TEMPF=DATAI(I)	0388
	TEMPJ=DATAI(J)	0389
	GO TO 403	0390
402	TEMPF=DATAI(I)	0391
	TEMPJ=DATAI(J)	0392
403	DATAI(I)=DATAI(J)-TEMPF	0393
	DATAI(I)=DATAI(J)-TEMPF	0394
	DATAI(J)=DATAI(J)+TEMPF	0395
	DATAI(J)=DATAI(J)+TEMPF	0396
410	J=J+ISTEP	0397
	IMIN=IMIN+IMIN-I1	0398
	ISTEP=ISTEP+ISTEP	0399

420	IF(ISTEP-NTWO)400,400,430	0400
430	CONTINUE	0401
C		0402
C	MAIN LOOP FOR FACTORS OF TWO. $W=EXP(1-2*PI*SQRT(-1)*M/MAX)$	0403
C		0404
	THETA=-THETA/8.	0405
	WSTEP=0.	0406
	WSTFI=-1.	0407
	IF(1FFWD)502,501,502	0408
501	THETA=-THETA	0409
	WSTFI=1.	0410
502	MMAX=A*NF1	0411
	GO TO 540	0412
500	WMIN=COS(THETA)	0413
	WMINI=SIN(THETA)	0414
	WR=WMIN	0415
	WI=WMINI	0416
	MMIN=M*AX/2+NF1	0417
	MSTEP=NF1*NF1	0418
	DO 530 M=MMIN,MMAX,MSTEP	0419
	DO 525 I=1,IIRNG	0420
	ISTEP=M*AX	0421
	IMIN=M+I1	0422
510	J=IMIN-ISTEP/2	0423
	DO 520 I=IMIN,NTOT,ISTEP	0424
	TEMPF=DATAI(I)*WR-DATAI(J)*WI	0425
	TEMPJ=DATAI(I)*WI+DATAI(J)*WR	0426
	DATAI(I)=DATAI(J)-TEMPF	0427
	DATAI(J)=DATAI(I)+TEMPJ	0428
	DATAI(I)=DATAI(I)+TEMPF	0429
	DATAI(J)=DATAI(J)+TEMPJ	0430
520	J=J+ISTEP	0431
	IMIN=IMIN+IMIN-I1	0432
	ISTEP=ISTEP+ISTEP	0433
	IF(ISTEP-NTWO)510,510,525	0434
525	CONTINUE	0435
	WTEMP=WR*WSTFI	0436
	WR=WR*WSTEP-WI*WSTFI	0437
530	WI=WI*WSTEP+WTEMP	0438
	WSTEP=WMIN	0439
	WSTFI=WMINI	0440
	THETA=THETA/2.	0441
	MMAX=M*AX+MMAX	0442
540	IF(MMAX-NTWO)500,500,600	0443
C		0444
C	MAIN LOOP FOR FACTORS NOT EQUAL TO TWO.	0445
C	$W=EXP(1-2*PI*SQRT(-1)*(J2-I3)/IFF2)$	0446
C		0447
600	IF(INGN2F-1)700,700,601	0448
601	IFF1=NTWO	0449

	IF=IACN2	0450
610	IFF2=IFACT(IFI)*IFF1	0451
	THETA=-THOFI/FLCAT(IFACT(IFI))	0452
	IF(1/IFFND)612,611,612	0453
611	THETA=-THETA	0454
612	THETM=THETA/FLCAT(IFF1/NF1)	0455
	WSTPR=COS(THETA)	0456
	WSTPI=SIN(THETA)	0457
	WMSTR=COS(THETM)	0458
	WMSTI=SIN(THETM)	0459
	WMINF=1.	0460
	WMINI=C.	0461
	DO 660 J1=1,IFF1,NF1	0462
	I1MAX=J1+I1FAG-1	0463
	DO 650 I1=J1,I1MAX	0464
	DO 650 I3=I1,NICT,AP2	0465
	I=1	0466
	WR=WMINF	0467
	WI=WMINI	0468
	J2MAX=I3+IFF2-IFF1	0469
	DO 640 J2=I3,J2MAX,IFF1	0470
	TWOWF=WR*W	0471
	JMIN=I3	0472
	J3MAX=J2+NF2-IFF2	0473
	DO 630 J3=J2,J3MAX,IFF2	0474
	J=JMIN+IFF2-IFF1	0475
	SR=DATAR(J)	0476
	SI=DATAI(J)	0477
	OLDSF=0.	0478
	OLDSI=0.	0479
	J=J-IFF1	0480
620	STMPR=SF	0481
	STMDI=SI	0482
	SF=TWOWF*SF-OLDSR+DATAR(J)	0483
	SI=TWOWF*SI-OLDSI+DATAI(J)	0484
	OLDSF=STMPR	0485
	OLDSI=STMDI	0486
	J=J-IFF1	0487
	IF (J-JMIN)621,621,620	0488
621	WOPKP(I)-WR*SF-WI*SI-OLDSR+DATAR(J)	0489
	WOPKI(I)-WI*SR+WR*SI-OLDSI+DATAI(J)	0490
	JMIN=JMIN+IFF2	0491
630	I=I+1	0492
	WTEMP=WR*WSTPI	0493
	WR=WR*WSTPR-WI*WSTPI	0494
640	WI=WI*WSTPR+WTEMP	0495
	I=1	0496
	DO 650 J2=I3,J2MAX,IFF1	0497
	J3MAX=J2+NF2-IFF2	0498
	DO 650 J3=J2,J3MAX,IFF2	0499

	CATAR(J3)=WORKF(I)	C500
	DATAI(J3)=WCRKI(I)	C501
650	I=I+1	C502
	WTEMP=WMINR*WMSTI	C503
	WMINR=WMINR*WMSTP-WMINI*WMSTI	C504
660	WMINI=WMINI*WMSTR+WTEMP	C505
	IF=IF+1	C506
	IFF1=IFF2	C507
	IF(IFF1-NP2)610,700,700	C508
C		C509
C	COMPLETE A REAL TRANSFORM IN THE 1ST DIMENSION, A EVEN, BY CON-	C510
C	JUGATE SYMMETRIES.	C511
C		C512
700	GO TO (500,800,900,701),ICASE	C513
701	NHALF=N	C514
	N=N+N	C515
	THETA=-TWOPI/FLOAT(N)	C516
	IF(IFFWD)703,702,703	C517
702	THETA=-THETA	C518
703	WSTPR=(COS(THETA)	C519
	WSTPI=SIN(THETA)	C520
	WF=WSTPR	C521
	WI=WSTPI	C522
	IMIN=2	C523
	JMIN=NHAF	C524
	GO TO 725	C525
710	J=JMIN	C526
	DO 720 I=IMIN,NTOT,NP2	C527
	SUMP=(CATAR(I)+CATAR(J))/2.	C528
	SUMI=(DATAI(I)+DATAI(J))/2.	C529
	DIFF=(CATAR(I)-CATAR(J))/2.	C530
	DIFI=(DATAI(I)-DATAI(J))/2.	C531
	TEMPR=WR*SUMI+WI*DIFF	C532
	TEMPI=WI*SUMI-WR*DIFF	C533
	CATAR(I)=SUMP+TEMP	C534
	DATAI(I)=DIFI+TEMPI	C535
	CATAR(J)=SUMP-TEMP	C536
	DATAI(J)=-DIFI+TEMPI	C537
720	J=J+NP2	C538
	IMIN=I+IN+1	C539
	JMIN=JMIN+1	C540
	WTEMP=WR*WSTPI	C541
	WR=WR*WSTPR-WI*WSTPI	C542
	WI=WI*WSTPR+WTEMP	C543
725	IF(I+IN-JMIN)710,710,740	C544
730	IF(IFFWD)731,740,731	C545
731	DO 735 I=I+IN,NTOT,NP2	C546
735	CATAR(I)=-CATAR(I)	C547
740	NP2=NP2+NP2	C548
	NTOT=NTOT+NTOT	C549

	J=NTOT+1	0550
	IMAX=NTOT/2+1	0551
745	IMIN=IMAX-NHALF	0552
	I=IMIN	0553
	GO TO 755	0554
750	DATA(I)=DATA(J)	0555
	DATA(J)=-DATA(I)	0556
755	I=I+1	0557
	J=J-1	0558
	IF(I-IMAX)750,760,760	0559
760	DATA(J)=DATA(IMIN)-DATA(IMIN)	0560
	DATA(J)=0.	0561
	IF(I-J)770,760,780	0562
765	DATA(J)=DATA(I)	0563
	DATA(J)=DATA(I)	0564
770	I=I-1	0565
	J=J-1	0566
	IF(I-IMIN)775,775,765	0567
775	DATA(J)=DATA(IMIN)+DATA(IMIN)	0568
	DATA(J)=0.	0569
	IMAX=IMIN	0570
	GO TO 745	0571
780	DATA(I)=DATA(I)+DATA(I)	0572
	DATA(I)=0.	0573
	GO TO 900	0574
C		0575
C	COMPLETE A REAL TRANSFORM FOR THE 2ND, 3RD, ETC. DIMENSION BY	0576
C	CONJUGATE SYMMETRIES.	0577
C		0578
800	IF(NP2V-2)900,900,805	0579
805	GO 860 I3=1,NTOT,NF2	0580
	I2MAX=I3+NF2-NF1	0581
	GO 860 I2=I3,I2MAX,NF1	0582
	IMAX=I2+NP1-1	0583
	IMIN=I2+1	0584
	JMAX=I3+I2+NP1-IMIN	0585
	IF(I2-I3)820,820,810	0586
810	JMAX=JMAX+NF2	0587
820	IF(I0IM-2)850,850,830	0588
830	J=JMAX+NF0	0589
	DO 840 I=IMIN,IMAX	0590
	DATA(I)=DATA(J)	0591
	DATA(I)=-DATA(J)	0592
840	J=J-1	0593
850	J=JMAX	0594
	DO 860 I=IMIN,IMAX,NF0	0595
	DATA(I)=DATA(J)	0596
	DATA(I)=-DATA(J)	0597
860	J=J-NF0	0598
C		0599

C END OF LOOP ON EACH DIMENSION
C
900 NPG=NF1
 NP1=NF2
910 NPREV=N
920 RETURN
 END

0600
0601
0602
0603
0604
0605
0606



DATE
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